New statistical method for enhancing the real-time recognition of astrophysical neutrino burst

PhD Candidate Marco Mattiazzi

Università di Siena, and INFN Padova

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Real-time detection of astrophysical ν burst

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Overview



- 2 Clustering algorithms
- 3 Standard fixed-threshold method

4 Beta filter method

- Dynamic clustering
- Shifted clustering
- 5 Performance Evaluation
 - 6 Conclusions

Main objective of this seminar

Aim

Enhance the **real-time** recognition of astrophysical neutrino burst in a **statistical** fashion

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Requirements

- **short-latency** \longrightarrow prompt alert to ν , EM and GW telescopes
- low false alarm rate crucial in the multi-messenger era

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How-to

Exploit the different **time structure** of the **signal** with respect to the (*poissonian*) **background**

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Application : ν -driven Core-Collapse Supernovae

CCSNe

- last stage lifecycle of massive stars
- released energy $\sim 10^{53} \text{ erg} \longrightarrow 99\%$ emitted as neutrinos
- models based on highly complex simulations but **exact mechanism is still unknown**
- so far, only ν -bursts from SN1987A



Potential scientific outcomes from the next CCSN neutrino detection

- neutrino mass limit + mass hierarchy + constraints on BSM physics
- unique **probe** for black-hole forming SNe or other *exotic* explosions
- **Real-time detection**

guide the optical telescopes to study EM signal from its onset

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Real-time detection

• guide the optical telescopes to study EM signal from its onset

Clustering algorithms

Dynamic fixed-size sliding time windows start from each event



Background Signal

Static fixed-size sliding time windows without overlapping



Shifted same as static + second translated scan to reduce search bias



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The *standard method* for the real-time **identification** of **candidate SN cluster** is based on the so-called **imitation frequency**

$$egin{aligned} \mathrm{F}_{im}(m \mid r, w) &= \mathit{N}_{\mathrm{windows}}(t_{FAT}) imes \mathit{P}_{\mathrm{survival}}(m \mid rw) \ &= \mathit{N}_{\mathrm{windows}}(t_{FAT}) imes \sum_{k=m}^{\infty} rac{(rw)^k e^{-rw}}{k!} \end{aligned}$$

and the **fixed-threshold** \hat{M} is defined such that

 $\mathrm{F}_{\mathit{im}}(\hat{M} \mid r, w) < 1$

where:

- $w \longrightarrow \text{time window size } (e.g. 20 \text{ s})$
- $t_{FAT} \longrightarrow$ required false alert time (e.g. 1 century)
- $r \longrightarrow$ Poisson background rate
- $m \longrightarrow$ multiplicity, *i.e.* number of events within the analysed time window

Standard method based on Imitation Frequency

Main Issue

Cut only on $\textit{multiplicity} \longrightarrow$ no sensitivity on the time structure of the SN burst





 Δt duration of the cluster, *i.e.* time difference between the first and the last event

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The statistical basis behind the newly proposed β method

Hypothesis: stationary Poisson process with constant rate r

$$P_{\text{survival}}(m \mid rw) = 1 - \sum_{k=0}^{m-1} \frac{(rt)^k e^{-rt}}{k!} = \sum_{k=m}^{\infty} \text{Pois}(k, rt)$$

probability of observing at least m events in a given time interval [0, t]

Key Idea: Order Statistics

In a fixed time window W = [0, w], the *m* events are uniformly distributed with

$$0 < t_1 < t_2 < \cdots < t_m < w$$

Normalising to [0,1] via

 $x_k := t_k/w$

then the normalised time of the k-th event follows $f(x_k \mid m, w) = \text{Beta}(x_k; k, m+1-k)$ *i.e.* the **k-th order statistic** of **uniform distribution**

Transforming back in time domain

$$f(t_k \mid m, w) = \frac{1}{w} \times \text{Beta}\left(\frac{t_k}{w}; k, m+1-k\right)$$

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Definition

Let X_1, X_2, \dots, X_N be N random variables *i.i.d* with F(x) distribution, sorting in increasing order

$$X_{(1)} \leq X_{(2)} \leq \cdots \leq X_{(N)}$$

are r.v. as well, called order statistics

Cumulative distribution

Iargest order statistic X_(N) = max{X₁, X₂, · · · , X_N}

$$F_{X_{(N)}}(x) = \Pr(X_{(N)} \le x) = [F(x)]'$$

• first order statistic $X_{(1)} = \min\{X_1, X_2, \cdots, X_N\}$

$$F_{X_{(1)}}(x) = \Pr(X_{(1)} \le x) = 1 - [1 - F(x)]^n$$

k-th order statistic

$$F_{X_{(k)}}(x)={\sf Pr}(ext{ at least }k ext{ of }X ext{ 's are }\leq x)=\sum_{i=1}^n \binom{n}{i}\,[F(x)]^j\,[1-F(x)]^{n-j}$$

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Cumulative distribution

• largest order statistic $X_{(N)} = \max\{X_1, X_2, \cdots, X_N\}$

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PDF k-th order statistic

$$f_{X_{(k)}}(x) = \frac{n!}{(k-1)!(n-k)!} f(x) [F(x)]^{k-1} [1-F(x)]^{n-k}$$



Heuristic Proof

 $\begin{aligned} \Pr(X_{(k)} \in [x, x + dx]) &= \Pr(\text{one of the } X' \text{s} \in [x, x + dx] \quad \text{and} \quad k - 1 \quad \text{of the others } X' \text{s} < x) \\ &= [n \Pr(X \in [x, x + dx])] \Pr(k - 1 \text{ of other } X' \text{s} < x) \\ &= n f(x) \, dx \quad \binom{n - 1}{k - 1} [\Pr(X < x)]^{k - 1} [\Pr(X > x)]^{n - k} \\ &= n \binom{n - 1}{k - 1} f(x) \, dx \quad [F(x)]^{k - 1} [1 - F(x)]^{n - k} \end{aligned}$

Recalling Gamma function definition for integers, $\Gamma(n+1) = n!$

$$f_{X_{(k)}}(x) = \frac{\Gamma(n+1)}{\Gamma(n-k+1)\Gamma(k)} f(x)[F(x)]^{k-1} [1-F(x)]^{n-k}$$

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Beta distribution

$$\mathsf{Beta}(u;\alpha,\beta) = \frac{u^{\alpha-1}(1-u)^{\beta-1}}{\int_0^1 v^{\alpha-1}(1-v)^{\beta-1} \, dv} = \frac{\Gamma(\alpha+\beta)}{\Gamma(\alpha)\Gamma(\beta)} \, u^{\alpha-1}(1-u)^{\beta-1} \qquad u \in [0,1]$$

Order statistics of Uniform distribution on Unit interval

Let X_1, X_2, \dots, X_N be N random variables *i.i.d* with $X \sim \text{Uniform}(0,1)$, the k-th order statistic

$$f_{X_{(k)}}(u) = \frac{n!}{(k-1)!(n-k)!} u^{k-1} (1-u)^{n-k} \quad \text{with } 0 \le u \le 1$$
$$= \frac{\Gamma(n+1)}{\Gamma(n-k+1)\Gamma(k)} u^{k-1} (1-u)^{n-k}$$
$$= \text{Beta}(u; k, n-k+1)$$

Thus

$$X_{(k)} \sim \operatorname{Beta}(k, n-k+1)$$

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Thus

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Dynamic clustering

PDF: by construction
$$t_1 = 0 \longrightarrow \Delta t = t_m$$

 $f_D(\Delta t \mid m, w) = f_\beta(t_m \mid m - 1, w)$
 $= \frac{1}{w} \times \text{Beta}\left(\frac{\Delta t}{w}; m - 1, 1\right)$
 $= \frac{m - 1}{w^{m-1}} \times \Delta t^{m-2}$

CDF: fixed multiplicity *m*

$$egin{aligned} \mathcal{F}_{\mathrm{D}}(\Delta t \mid m) &= \mathrm{Prob}(t \leq \Delta t \mid m, w) \ &= \int_{0}^{\Delta t} dt \, f_{\mathrm{D}}(t \mid m, w) \ &= rac{1}{m} \mathrm{Beta}\left(rac{\Delta t}{w}; m, 1
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Numerical results + theoretical distributions

Dynamic clustering

Discrete multiplicity density distribution \longrightarrow truncated Poisson pdf

$$g_{\rm D}(m \mid r, w) = \frac{(rw)^{m-1} e^{-rw}}{(m-1)!} \frac{1}{1 - e^{-rw}}$$
$$= \frac{{\rm Pois}(m-1, \mu = rw)}{1 - e^{-\mu}}$$

Joint probability density distribution

$$j_{\mathrm{D}}(m,\Delta t \mid r,w) = f_{\mathrm{D}}(\Delta t \mid m,w) \times g_{\mathrm{D}}(m \mid r,w)$$
$$= \frac{re^{-rw}}{1 - e^{-rw}} \frac{(r\Delta t)^{m-2}}{(m-2)!}$$

Key Idea: Decision Boundary

Build decision boundary, $\{\Delta t_k\}$ with $k \ge 2$, on the $m - \Delta t$ plane such that the required false alert rate is ensured in analytical way \longrightarrow no MC-based thresholds

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Dynamic clustering

 $\textbf{Discrete multiplicity density distribution} \longrightarrow \texttt{truncated Poisson pdf}$

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Dynamic Clustering

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Constraint

$$\sum_{k=2}^{\infty} \mathit{N}_k imes \mathit{F}_{\mathrm{D}}(\Delta t_k \mid k) \leq 1$$

Expected number of clusters with multiplicity k over t_{FAT}

$$N_k = rt_{FAT} \times \mathrm{Pois}(k-1, \mu = rw)$$

Find \hat{M} such that

$$\sum_{j=\hat{M}}^{\infty} N_j < 1$$

e.g. the standard imitation frequency threshold

Set $\Delta t_j = w$ for $j \ge \hat{M}$, the constraint can be rearranged as

$$\sum_{k=2}^{\hat{M}-1} \operatorname{Pois}\left(k-1,\mu\right) \times \left(\frac{\Delta t_k}{w}\right)^{k-1} \leq \frac{1-\sum_{j=\hat{M}}^{\infty} N_j}{r t_{FAT}}$$

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β discrete filter

Introducing the $\{\beta_k\}_{k=2}^{\hat{M}-1}$ discrete filter, as follows

$$\operatorname{Pois}(k-1,\mu) \times \left(\frac{\Delta t_k}{w}\right)^{k-1} = \beta_k \times \frac{1 - \sum_{j=\hat{M}}^{\infty} N_j}{r t_{FAT}}$$
$$\sum_{k=2}^{\hat{M}-1} \beta_k = 1$$

Such set of equations allow exploring the region $m < \hat{M}$

Decision boundary : Analytical Formula

Finally we get

$$\Delta t_k = \begin{cases} w \left[\beta_k \cdot \frac{1 - \sum_{j=\hat{M}}^{\infty} N_j}{N_k} \right]^{1/(k-1)} & \text{if } k < \hat{M} \\ w & \text{if } k \ge \hat{M} \end{cases}$$

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Dynamic clustering: β filter application

Example: dynamic clustering, w =20 s, t_{FAT} = 1 day, r = 0.1 Hz, N_{simu} = $10^4 \times t_{FAT}$

Remark on β filter method

High flexibility of β filter method \longrightarrow set additional constraints based on e.g.

- penalising region where non-poissonian background occurs more frequently
- discarding the multiplicities with Δt_k below the timing resolution of the detector

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Shifted clustering

PDF: $\Delta t = t_m - t_1 \longrightarrow$ **sample range** of order statistics of uniform distribution

$$f_{\rm S}(\Delta t|m,w) = \frac{m(m-1)}{w} \left(\frac{\Delta t}{w}\right)^{m-2} \left(1 - \frac{\Delta t}{w}\right)$$
$$= \frac{1}{w} \times \text{Beta}\left(\frac{\Delta t}{w}; m-1, 2\right)$$

CDF: fixed multiplicity *m*

$$\begin{split} \overline{F}_{\mathrm{S}}(\Delta t \mid m) &= \mathrm{Prob}(t \leq \Delta t \mid m, w) \\ &= \int_{0}^{\Delta t} dt \, f_{\mathrm{S}}(t \mid m, w) \\ &= \mathrm{Beta}\left(\frac{\Delta t}{w}; m, 1\right) \times \\ &\left[1 - \left(\frac{m-1}{m}\right)\left(\frac{\Delta t}{w}\right)\right] \end{split}$$

Shifted clustering

discrete multiplicity pdf

$$g_{\rm S}(m \mid r, w) = \frac{\operatorname{Pois}(m, rw)}{1 - e^{-rw}(1 + rw)}$$

• # clusters, multiplicity k over t_{FAT}

$$N_k = N_W \times \text{Pois}(k, \mu = rw)$$

where $N_W = (2t_{FAT})/w - 1$

Likewise the dynamic case

$$N_k imes F_{
m S}(\Delta t_k \,|\, k) = eta_k imes [1 - \sum_{j=\hat{M}}^\infty N_j] \quad ext{for } k < \hat{M}$$

solvable for $\{\Delta t_k\}_{k=2}^{\tilde{M}-1} \longrightarrow$ root between [0, w] of the associated k-th degree polynomials

Shifted clustering

• discrete multiplicity pdf $g_{\rm S}(m \mid r, w) = \frac{\text{Pois}(m, rw)}{1 - e^{-rw}(1 + rw)}$ • # clusters, multiplicity k over t_{FAT} $N_k = N_W \times \text{Pois}(k, \mu = rw)$ where $N_W = (2t_{FAT})/w - 1$

Likewise the dynamic case

$$egin{aligned} \mathsf{N}_k imes \mathsf{F}_{\mathrm{S}}(\Delta t_k \,|\, k) &= eta_k imes [1\!-\!\sum_{j=\hat{M}}^\infty \mathsf{N}_j] \quad ext{for } k < \hat{M} \end{aligned}$$

solvable for $\{\Delta t_k\}_{k=2}^{\hat{M}-1} \longrightarrow$ root between [0, w] of the associated k-th degree polynomials

Performance Enhancement Evaluation

Shifted clustering, r =0.03 Hz, w =20 s, $au_{
m short}$ = 10 ms and $au_{
m long}$ = 1 s

SN-like ν signal

$$f_{
m signal}(t) = e^{-rac{t}{ au_{
m long}}} (1-e^{-rac{t}{ au_{
m short}}})$$

Performance Enhancement Evaluation

Dynamic clustering, w = 20 s, $\tau_{\rm short} = 10$ ms and $\tau_{\rm long} = 1$ s

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Summary

New β filter method

Main Result

• important performance improvement (up to \sim 80%) for real-time SN- ν monitor with respect to standard technique

Pro's

- exploit timing information \longrightarrow bi-dimensional cut on $m \Delta t$ plane
- ullet no MC-based thresholds \longrightarrow only 1D root-finding algorithm required
- flexibility in the construction of {∆t_k}^{M−1}_{k=2} decision boundary → allow setting detector-based thresholds
- independent on the nature of clusters applicable on wide class of real-time discrimination processes where sliding-time windows are involved + fixed false discovery rate is required

Con's

• careful identification of non-poissonian background component

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Backup

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[Backup] Joint PDF of two order statistics

Likewise the pdf of k-th order statistic's case

$$\begin{aligned} &f_{X_{(j)},X_{(k)}}(x,y)dx\,dy = n\,(n-1) & \cdots & f(x)f(y)\,dx\,dy \quad \text{where } x \leq y \\ &f_{X_{(j)},X_{(k)}}(x,y) = \frac{n(n-1)(n-2)!}{(j-1)!(k-j-1)!(n-k)!}[F(x)]^{j-1}[F(y)-F(x)]^{k-1-j}[1-F(y)]^{n-k}f(x)f(y) \end{aligned}$$

$$f_{X_{(j)},X_{(k)}}(x,y) = \frac{n!}{(j-1)!(k-j-1)!(n-k)!} [F(x)]^{j-1} [F(y) - F(x)]^{k-1-j} [1 - F(y)]^{n-k} f(x) f(y)$$

[Backup] Sample range of o.s. of uniform distribution

$$f_{X_{(1)},X_{(n)}}(x,y) = \frac{n!}{(n-2)!} [F(y) - F(x)]^{n-2} f(x) f(y)$$

Let X_1, X_2, \dots, X_N be N random variables *i.i.d* with $X \sim \text{Uniform}(0, 1)$

$$f_{X_{(1)},X_{(n)}}(x,y) = n(n-1)[y-x]^{n-2}$$
 where $0 \le x \le y \le 1$

change of variable w = y - x and z = x

$$f_{X_{(1)},X_{(n)}}(w,z) = n(n-1)w^{n-2}$$

marginalising, $y \leq 1 \longrightarrow w + z \leq 1$

$$f_{X_{(1)},X_{(n)}}(w) = \int_0^{1-w} f_{X_{(1)},X_{(n)}}(w,z)dz$$

= $n(n-1)w^{n-2}(1-w)$
= $Beta(w; n-1,2)$

Thus sample range of order statistics of Uniform distribution

$$f_{1,n}(w) \sim \mathsf{Beta}(w; n-1, 2)$$

Shifted clustering, r = 0.03 Hz, w = 20 s, $N_{\text{simulations}} = 10^7$

Dynamic Uniform clustering, w =20 s, $au_{
m short}$ = 10 ms and $au_{
m long}$ = 1 s

Dynamic Ramp clustering, w =20 s, $au_{
m short}$ = 10 ms and $au_{
m long}$ = 1 s

Shifted Uniform clustering, w =20 s, $au_{
m short} = 10$ ms and $au_{
m long} = 1$ s

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Shifted Ramp clustering, w =20 s, $au_{
m short} = 10$ ms and $au_{
m long} = 1$ s

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