# <span id="page-0-0"></span>New statistical method for enhancing the real-time recognition of astrophysical neutrino burst

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Universit`a di Siena, and INFN Padova

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# <span id="page-1-0"></span>**Overview**



- 2 [Clustering algorithms](#page-7-0)
- 3 [Standard fixed-threshold method](#page-10-0)

### 4 [Beta filter method](#page-12-0)

- [Dynamic clustering](#page-21-0)
- [Shifted clustering](#page-32-0)

## 5 [Performance Evaluation](#page-35-0)

## **[Conclusions](#page-37-0)**

# <span id="page-2-0"></span>Main objective of this seminar

### Aim

Enhance the real-time recognition of astrophysical neutrino burst in a statistical fashion

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# Main objective of this seminar

### Aim

Enhance the real-time recognition of astrophysical neutrino burst in a statistical fashion

### **Magazine (1999)** Requirements

- short-latency  $\longrightarrow$  prompt alert to  $\nu$ , EM and GW telescopes
- **low** false alarm rate  $\rightarrow$  crucial in the multi-messenger era

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# <span id="page-4-0"></span>Main objective of this seminar

### Aim

Enhance the real-time recognition of astrophysical neutrino burst in a statistical fashion

### No principal de Mar Requirements

- short-latency  $\longrightarrow$  prompt alert to  $\nu$ , EM and GW telescopes
- **low** false alarm rate  $\rightarrow$  crucial in the multi-messenger era

a nacional code How-to

Exploit the different **time structure** of the **signal** with respect to the (poissonian) background

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# <span id="page-5-0"></span>Application :  $\nu$ -driven Core-Collapse Supernovae

### CCSNe

- **last stage lifecycle of massive stars**
- released energy  $\sim 10^{53}$  erg  $\longrightarrow 99\%$ emitted as neutrinos
- models based on highly complex simulations but exact mechanism is still unknown
- $\bullet$  so far, only  $\nu$ -bursts from SN1987A



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- - **•** guide the [o](#page-4-0)p[t](#page-7-0)ical telescopes to study **EM sig[n](#page-5-0)al** fro[m](#page-4-0) i[ts](#page-6-0) **on[se](#page-6-0)t**

# <span id="page-6-0"></span>Application : ν-driven Core-Collapse Supernovae

### CCSNe

- **last stage lifecycle of massive stars**
- released energy  $\sim 10^{53}$  erg  $\longrightarrow 99\%$ emitted as neutrinos
- models based on highly complex simulations but exact mechanism is still unknown
- $\bullet$  so far, only  $\nu$ -bursts from SN1987A



Potential scientific outcomes from the next CCSN neutrino detection

- neutrino mass limit  $+$  mass hierarchy  $+$  constraints on BSM physics
- **.** unique probe for black-hole forming SNe or other exotic explosions

### Real-time detection

**e** guide the [o](#page-4-0)p[t](#page-7-0)ical telescopes to study **EM sig[n](#page-5-0)al** fro[m](#page-5-0) i[ts](#page-7-0) **on[se](#page-6-0)t** 

# <span id="page-7-0"></span>Clustering algorithms

Dynamic fixed-size sliding time windows start from each event



### **Background Signal**





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# Clustering algorithms

Dynamic fixed-size sliding time windows start from each event



### **Background Signal**

Static fixed-size sliding time windows without overlapping





 $\leftarrow$   $\Box$ 

 $QQQ$ 

# Clustering algorithms

Dynamic fixed-size sliding time windows start from each event



### **Background Signal**

Static fixed-size sliding time windows without overlapping



**Shifted** same as static  $+$  second translated scan to reduce search bias



 $QQQ$ 

<span id="page-10-0"></span>The *standard method* for the real-time **identification** of **candidate SN cluster** is based on the so-called imitation frequency

$$
F_{im}(m \mid r, w) = N_{\text{windows}}(t_{FAT}) \times P_{\text{survival}}(m \mid rw)
$$

$$
= N_{\text{windows}}(t_{FAT}) \times \sum_{k=m}^{\infty} \frac{(rw)^k e^{-rw}}{k!}
$$

and the fixed-threshold  $\hat{M}$  is defined such that

 $F_{im}(\hat{M} \mid r, w) < 1$ 

where:

- $w \rightarrow$  time window size (e.g. 20 s)
- $t_{FAT} \longrightarrow$  required false alert time (e.g. 1 century)
- $\bullet$   $r \rightarrow$  Poisson background rate
- $\bullet$  m  $\rightarrow$  multiplicity, *i.e.* number of events within the analysed time window

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# <span id="page-11-0"></span>Standard method based on Imitation Frequency

### Main Issue

Cut only on *multiplicity*  $\longrightarrow$  no sensitivity on the time structure of the SN burst





∆t duration of the cluster, i.e. time difference between the first and the last event

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# <span id="page-12-0"></span>The statistical basis behind the newly proposed  $\beta$  method

Hypothesis: stationary Poisson process with constant rate r

$$
P_{\text{survival}}(m \mid rw) = 1 - \sum_{k=0}^{m-1} \frac{(rt)^k e^{-rt}}{k!} = \sum_{k=m}^{\infty} \text{Pois}(k, rt)
$$

**probability** of observing at least m events in a given time interval  $[0, t]$ 

$$
0 < t_1 < t_2 < \cdots < t_m < w
$$

$$
f(t_k \mid m, w) = \frac{1}{w} \times \text{Beta}\left(\frac{t_k}{w}; k, m+1-k\right)
$$

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**probability** of observing at least m events in a given time interval  $[0, t]$ 

Key Idea: Order Statistics

In a fixed time window  $W = [0, w]$ , the m events are uniformly distributed with

$$
0
$$

Normalising to [0,1] via

 $x_k := t_k/w$ 

then the normalised time of the k-th event follows  $f(x_k | m, w) = \text{Beta}(x_k; k, m + 1 - k)$ i.e. the k-th order statistic of uniform distribution

$$
f(t_k \mid m, w) = \frac{1}{w} \times \text{Beta}\left(\frac{t_k}{w}; k, m+1-k\right)
$$

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# <span id="page-14-0"></span>The statistical basis behind the newly proposed  $\beta$  method

Hypothesis: stationary Poisson process with constant rate r

$$
P_{\text{survival}}(m \mid rw) = 1 - \sum_{k=0}^{m-1} \frac{(rt)^k e^{-rt}}{k!} = \sum_{k=m}^{\infty} \text{Pois}(k, rt)
$$

**probability** of observing at least m events in a given time interval  $[0, t]$ 

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Transforming back in time domain

$$
f(t_k \mid m, w) = \frac{1}{w} \times \text{Beta}\left(\frac{t_k}{w}; k, m+1-k\right)
$$

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### <span id="page-15-0"></span>Definition

Let  $X_1, X_2, \cdots, X_N$  be N random variables *i.i.d* with  $F(x)$  distribution, sorting in increasing order

$$
X_{(1)}\leq X_{(2)}\leq\cdots\leq X_{(N)}
$$

are r.v. as well, called order statistics

**largest order statistic**  $X_{(N)} = \max\{X_1, X_2, \cdots, X_N\}$ 

$$
F_{X_{(N)}}(x) = \Pr(X_{(N)} \le x) = [F(x)]'
$$

**•** first order statistic  $X_{(1)} = \min\{X_1, X_2, \cdots, X_N\}$ 

$$
F_{X_{(1)}}(x) = \Pr(X_{(1)} \le x) = 1 - [1 - F(x)]^n
$$

### Definition

Let  $X_1, X_2, \cdots, X_N$  be N random variables *i.i.d* with  $F(x)$  distribution, sorting in increasing order

$$
X_{(1)}\leq X_{(2)}\leq\cdots\leq X_{(N)}
$$

are r.v. as well, called order statistics

### Cumulative distribution

**largest order statistic**  $X_{(N)} = \max\{X_1, X_2, \cdots, X_N\}$ 

$$
F_{X_{(N)}}(x) = \Pr(X_{(N)} \le x) = [F(x)]^n
$$

• first order statistic 
$$
X_{(1)} = \min\{X_1, X_2, \cdots, X_N\}
$$

$$
F_{X_{(1)}}(x) = \Pr(X_{(1)} \le x) = 1 - [1 - F(x)]^n
$$

**Q** k-th order statistic

$$
F_{X_{(k)}}(x) = \Pr(\text{ at least } k \text{ of } X\text{'s are } \leq x) = \sum_{j=k}^{n} {n \choose j} [F(x)]^{j} [1 - F(x)]^{n-j}
$$

### PDF k-th order statistic

$$
f_{X_{(k)}}(x) = \frac{n!}{(k-1)!(n-k)!} f(x) [F(x)]^{k-1} [1 - F(x)]^{n-k}
$$



$$
f_{X_{(k)}}(x) = \frac{\Gamma(n+1)}{\Gamma(n-k+1)\Gamma(k)} f(x)[F(x)]^{k-1}[1 - F(x)]^{n-k}
$$

### PDF k-th order statistic

$$
f_{X_{(k)}}(x) = \frac{n!}{(k-1)!(n-k)!} f(x) [F(x)]^{k-1} [1 - F(x)]^{n-k}
$$



### Heuristic Proof

$$
\Pr(X_{(k)} \in [x, x + dx]) = \Pr(\text{one of the } X's \in [x, x + dx] \text{ and } k - 1 \text{ of the others } X's < x)
$$
\n
$$
= [n \Pr(X \in [x, x + dx])] \Pr(k - 1 \text{ of other } X's < x)
$$
\n
$$
= n f(x) dx \binom{n - 1}{k - 1} [\Pr(X < x)]^{k - 1} [\Pr(X > x)]^{n - k}
$$
\n
$$
= n \binom{n - 1}{k - 1} f(x) dx \quad [F(x)]^{k - 1} [1 - F(x)]^{n - k}
$$

Recalling Gamma function definition for integers,  $\Gamma(n+1) = n!$ 

$$
f_{X_{(k)}}(x) = \frac{\Gamma(n+1)}{\Gamma(n-k+1)\Gamma(k)} f(x)[F(x)]^{k-1}[1 - F(x)]^{n-k}
$$

 $\Box$ 

### Beta distribution

Beta
$$
(u; \alpha, \beta)
$$
 =  $\frac{u^{\alpha-1}(1-u)^{\beta-1}}{\int_0^1 v^{\alpha-1}(1-v)^{\beta-1} dv}$  =  $\frac{\Gamma(\alpha+\beta)}{\Gamma(\alpha)\Gamma(\beta)} u^{\alpha-1}(1-u)^{\beta-1}$   $u \in [0,1]$ 

Let  $X_1, X_2, \dots, X_N$  be N random variables *i.i.d* with  $X \sim$  Uniform(0, 1), the k-th order statistic

$$
f_{X_{(k)}}(u) = \frac{n!}{(k-1)!(n-k)!}u^{k-1}(1-u)^{n-k} \quad \text{with } 0 \le u \le 1
$$

$$
= \frac{\Gamma(n+1)}{\Gamma(n-k+1)\Gamma(k)}u^{k-1}(1-u)^{n-k}
$$

$$
= \text{Beta}(u; k, n-k+1)
$$

$$
X_{(k)} \sim \textsf{Beta}(k,n-k+1)
$$

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### <span id="page-20-0"></span>Beta distribution

Beta
$$
(u; \alpha, \beta)
$$
 =  $\frac{u^{\alpha-1}(1-u)^{\beta-1}}{\int_0^1 v^{\alpha-1}(1-v)^{\beta-1} dv}$  =  $\frac{\Gamma(\alpha+\beta)}{\Gamma(\alpha)\Gamma(\beta)} u^{\alpha-1}(1-u)^{\beta-1}$   $u \in [0,1]$ 

### Order statistics of Uniform distribution on Unit interval

Let  $X_1, X_2, \cdots, X_N$  be N random variables *i.i.d* with  $X \sim$  Uniform(0, 1), the k-th order statistic

$$
f_{X_{(k)}}(u) = \frac{n!}{(k-1)!(n-k)!}u^{k-1}(1-u)^{n-k} \quad \text{with } 0 \le u \le 1
$$

$$
= \frac{\Gamma(n+1)}{\Gamma(n-k+1)\Gamma(k)}u^{k-1}(1-u)^{n-k}
$$

$$
= \text{Beta}(u; k, n-k+1)
$$

Thus

$$
X_{(k)} \sim \text{Beta}(k, n - k + 1)
$$

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## <span id="page-21-0"></span>Dynamic clustering

**PDF:** by construction 
$$
t_1 = 0 \longrightarrow \Delta t = t_m
$$
  
\n
$$
f_D(\Delta t \mid m, w) = f_\beta(t_m \mid m - 1, w)
$$
\n
$$
= \frac{1}{w} \times \text{Beta}\left(\frac{\Delta t}{w}; m - 1, 1\right)
$$
\n
$$
= \frac{m - 1}{w^{m - 1}} \times \Delta t^{m - 2}
$$

### CDF: fixed multiplicity m

$$
F_{D}(\Delta t \mid m) = \text{Prob}(t \le \Delta t \mid m, w)
$$

$$
= \int_{0}^{\Delta t} dt \, f_{D}(t \mid m, w)
$$

$$
= \frac{1}{m} \text{Beta}\left(\frac{\Delta t}{w}; m, 1\right)
$$

$$
= \left(\frac{\Delta t}{w}\right)^{m-1}
$$



Numerical results  $+$  theoretical distributions



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## Dynamic clustering

Discrete multiplicity density distribution → truncated Poisson pdf

$$
g_D(m \mid r, w) = \frac{(rw)^{m-1} e^{-rw}}{(m-1)!} \frac{1}{1 - e^{-rw}}
$$

$$
= \frac{\text{Pois}(m-1, \mu = rw)}{1 - e^{-\mu}}
$$

$$
j_{\rm D}(m,\Delta t \mid r, w) = f_{\rm D}(\Delta t \mid m, w) \times g_{\rm D}(m \mid r, w)
$$

$$
= \frac{re^{-rw}}{1 - e^{-rw}} \frac{(r\Delta t)^{m-2}}{(m-2)!}
$$

false alert rate is ensured in analytical way  $\longrightarrow$  no MC-based thresholds

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# <span id="page-23-0"></span>Dynamic clustering

Discrete multiplicity density distribution → truncated Poisson pdf

$$
g_D(m \mid r, w) = \frac{(rw)^{m-1} e^{-rw}}{(m-1)!} \frac{1}{1 - e^{-rw}}
$$

$$
= \frac{\text{Pois}(m-1, \mu = rw)}{1 - e^{-\mu}}
$$

Joint probability density distribution

$$
j_{\rm D}(m, \Delta t \mid r, w) = f_{\rm D}(\Delta t \mid m, w) \times g_{\rm D}(m \mid r, w)
$$
  
= 
$$
\frac{re^{-rw}}{1 - e^{-rw}} \frac{(r \Delta t)^{m-2}}{(m-2)!}
$$

### Key Idea: Decision Boundary

Build decision boundary,  $\{\Delta t_k\}$  with  $k \geq 2$ , on the  $m - \Delta t$  plane such that the required false alert rate is ensured in analytical way  $\longrightarrow$  no MC-based thresholds

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### *Dynamic Clustering*

<span id="page-24-0"></span>

### Key Idea: Decision Boundary

Build decision boundary,  $\{\Delta t_k\}$  with  $k \geq 2$ , on the  $m - \Delta t$  plane such that the required false alert rate is ensur[ed](#page-25-0) in ana[l](#page-31-0)ytical way  $\longrightarrow$  $\longrightarrow$  $\longrightarrow$  no MC-[ba](#page-23-0)sed [t](#page-23-0)[hr](#page-24-0)[e](#page-25-0)[sh](#page-20-0)ol[d](#page-32-0)[s](#page-11-0)

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### <span id="page-25-0"></span>**Constraint**

$$
\sum_{k=2}^\infty \mathsf{N}_k\times \mathsf{F}_{\mathrm{D}}(\Delta t_k\mid k)\leq 1
$$

Expected number of clusters with multiplicity  $k$  over  $t_{FAT}$ 

$$
N_k = rt_{FAT} \times \mathrm{Pois}(k-1, \mu = rw)
$$

$$
\sum_{j=\hat{M}}^{\infty}N_j<1
$$

$$
\sum_{k=2}^{\hat{M}-1} \mathrm{Pois}\left(k-1,\mu\right) \times \left(\frac{\Delta t_k}{w}\right)^{k-1} \leq \frac{1-\sum_{j=\hat{M}}^{\infty} N_j}{rt_{FAT}}
$$

### **Constraint**

$$
\sum_{k=2}^{\infty} N_k \times F_{\text{D}}(\Delta t_k \mid k) \leq 1
$$

Expected number of clusters with multiplicity  $k$  over  $t_{FAT}$ 

$$
N_k = rt_{FAT} \times \mathrm{Pois}(k-1, \mu = rw)
$$

Find  $\hat{M}$  such that

$$
\sum_{j=\hat{M}}^{\infty}N_j<1
$$

e.g. the standard imitation frequency threshold

$$
\sum_{k=2}^{M-1} \mathrm{Pois}\left(k-1, \mu\right) \times \left(\frac{\Delta t_k}{w}\right)^{k-1} \leq \frac{1-\sum_{j=\hat{M}}^{\infty} N_j}{r t_{FAT}}
$$

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### **Constraint**

$$
\sum_{k=2}^{\infty} N_k \times F_{\text{D}}(\Delta t_k \mid k) \leq 1
$$

Expected number of clusters with multiplicity  $k$  over  $t_{FAT}$ 

$$
N_k = rt_{FAT} \times \mathrm{Pois}(k-1, \mu = rw)
$$

Find  $\hat{M}$  such that

$$
\sum_{j=\hat{M}}^{\infty}N_j<1
$$

e.g. the standard imitation frequency threshold

Set  $\Delta t_i = w$  for  $j > \hat{M}$ , the constraint can be rearranged as

$$
\sum_{k=2}^{\hat{M}-1} \mathrm{Pois}\left(k-1,\mu\right) \times \left(\frac{\Delta t_k}{w}\right)^{k-1} \leq \frac{1-\sum_{j=\hat{M}}^{\infty} N_j}{r t_{FAT}}
$$

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### $\beta$  discrete filter

Introducing the  $\{\beta_k\}_{k=2}^{\hat{M}-1}$  discrete filter, as follows

$$
\text{Pois}(k-1,\mu) \times \left(\frac{\Delta t_k}{w}\right)^{k-1} = \beta_k \times \frac{1 - \sum_{j=\hat{M}}^{\infty} N_j}{r t_{\text{FAT}}}
$$

$$
\sum_{k=2}^{\hat{M}-1} \beta_k = 1
$$

Such set of equations allow exploring the region  $m < \hat{M}$ 

$$
\Delta t_k = \begin{cases} w \left[ \beta_k \cdot \frac{1 - \sum_{j=\hat{M}}^{\infty} N_j}{N_k} \right]^{1/(k-1)} & \text{if } k < \hat{M} \\ w & \text{if } k \geq \hat{M} \end{cases}
$$

### $\beta$  discrete filter

Introducing the  $\{\beta_k\}_{k=2}^{\hat{M}-1}$  discrete filter, as follows

Pois 
$$
(k-1, \mu) \times \left(\frac{\Delta t_k}{w}\right)^{k-1} = \beta_k \times \frac{1 - \sum_{j=\hat{M}}^{\infty} N_j}{r t_{\text{FAT}}}
$$

\n
$$
\sum_{k=2}^{\hat{M}-1} \beta_k = 1
$$

Such set of equations allow exploring the region  $m < \hat{M}$ 

### Decision boundary : Analytical Formula

Finally we get

$$
\Delta t_k = \begin{cases} w \left[ \beta_k \cdot \frac{1 - \sum_{j=\hat{M}}^{\infty} N_j}{N_k} \right]^{1/(k-1)} & \text{if } k < \hat{M} \\ w & \text{if } k \geq \hat{M} \end{cases}
$$

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## Dynamic clustering:  $\beta$  filter application



**Example:** dynamic clustering,  $w = 20$  s,  $t_{FAT} = 1$  day,  $r = 0.1$  Hz,  $N_{simu} = 10^4 \times t_{FAT}$ 



## <span id="page-31-0"></span>Remark on  $\beta$  filter method

High flexibility of  $\beta$  filter method  $\longrightarrow$  set additional constraints based on e.g.

- penalising region where non-poissonian background occurs more frequently
- $\bullet$  discarding the multiplicities with  $\Delta t_k$  below the timing resolution of the detector

 $\bullet$   $\cdots$ 



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## <span id="page-32-0"></span>Shifted clustering

PDF:  $\Delta t = t_m - t_1 \longrightarrow$  sample range of order statistics of uniform distribution

$$
f_{\rm S}(\Delta t | m, w) = \frac{m(m-1)}{w} \left(\frac{\Delta t}{w}\right)^{m-2} \left(1 - \frac{\Delta t}{w}\right)
$$

$$
= \frac{1}{w} \times \text{Beta}\left(\frac{\Delta t}{w}; m-1, 2\right)
$$

### CDF: fixed multiplicity m

$$
F_{\rm S}(\Delta t \mid m) = \text{Prob}(t \le \Delta t \mid m, w)
$$

$$
= \int_0^{\Delta t} dt \, f_{\rm S}(t \mid m, w)
$$

$$
= \text{Beta}\left(\frac{\Delta t}{w}; m, 1\right) \times \left[1 - \left(\frac{m - 1}{m}\right) \left(\frac{\Delta t}{w}\right)\right]
$$



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## <span id="page-33-0"></span>Shifted clustering

**o** discrete multiplicity pdf  $g_{\rm S}(m \mid r, w) = \frac{\text{Pois}(m, rw)}{1-e^{-rw}(1+rw)}$  $\bullet$  # clusters, multiplicity k over  $t_{FAT}$  $N_k = N_W \times \text{Pois}(k, \mu = rw)$ where  $N_W = (2t_{FAT})/w - 1$ 

$$
N_k \times F_{\rm S}(\Delta t_k \mid k) = \beta_k \times [1 - \sum_{j=\hat{M}}^{\infty} N_j] \quad \text{for } k < \hat{M}
$$



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## <span id="page-34-0"></span>Shifted clustering

**o** discrete multiplicity pdf  $g_{\rm S}(m \mid r, w) = \frac{\text{Pois}(m, rw)}{1-e^{-rw}(1+rw)}$  $\bullet$  # clusters, multiplicity k over  $t_{FAT}$  $N_k = N_W \times \text{Pois}(k, \mu = rw)$ where  $N_W = (2t_{FAT})/w - 1$ 

Likewise the dynamic case

$$
N_k \times F_{\mathrm{S}}(\Delta t_k \mid k) = \beta_k \times [1 - \sum_{j=\hat{M}}^{\infty} N_j] \quad \text{for } k < \hat{M}
$$

solvable for  $\left\{\Delta t_k\right\}_{k=2}^{\hat{M}-1} \longrightarrow$  root between  $[0,w]$ of the associated k-th degree polynomials



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## <span id="page-35-0"></span>Performance Enhancement Evaluation

**Shifted clustering**,  $r = 0.03$  Hz,  $w = 20$  s,  $\tau_{short} = 10$  ms and  $\tau_{long} = 1$  s



SN-like  $\nu$  signal

$$
f_{\text{signal}}(t) = e^{-\frac{t}{\tau_{\text{long}}}}(1-e^{-\frac{t}{\tau_{\text{short}}}})
$$

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## Performance Enhancement Evaluation

**Dynamic clustering**,  $w = 20$  s,  $\tau_{short} = 10$  ms and  $\tau_{long} = 1$  s



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# <span id="page-37-0"></span>Summary

### New  $\beta$  filter method

### Main Result

 $\bullet$  important performance improvement (up to  $\sim$  80%) for real-time SN- $\nu$  monitor with respect to standard technique

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# Summary

### New  $\beta$  filter method

### Main Result

 $\bullet$  important performance improvement (up to  $\sim$  80%) for real-time SN- $\nu$  monitor with respect to standard technique

### Pro's

- $\bullet$  exploit timing information  $\rightarrow$  bi-dimensional cut on  $m \Delta t$  plane
- no MC-based thresholds  $→$  only 1D root-finding algorithm required
- **flexibility** in the construction of  $\{\Delta t_k\}_{k=2}^{\hat{M}-1}$  decision boundary  $\longrightarrow$  allow setting detector-based thresholds
- independent on the nature of clusters  $\longrightarrow$  applicable on wide class of real-time discrimination processes where **sliding-time windows** are involved  $+$  fixed false discovery rate is required

### Con's

**•** careful identification of non-poissonian background component

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### <span id="page-41-0"></span>Backup

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# [Backup] Joint PDF of two order statistics



Likewise the pdf of k-th order statistic's case

$$
f_{X_{(j)},X_{(k)}}(x,y)dx dy = n(n-1) \qquad \cdots \qquad f(x)f(y) dx dy \quad \text{where } x \leq y
$$
  

$$
f_{X_{(j)},X_{(k)}}(x,y) = \frac{n(n-1)(n-2)!}{(j-1)!(k-j-1)!(n-k)!} [F(x)]^{j-1} [F(y) - F(x)]^{k-1-j} [1 - F(y)]^{n-k} f(x)f(y)
$$

$$
f_{X_{(j)},X_{(k)}}(x,y)=\frac{n!}{(j-1)!(k-j-1)!(n-k)!}[F(x)]^{j-1}[F(y)-F(x)]^{k-1-j}[1-F(y)]^{n-k}f(x)f(y)
$$

 $\leftarrow$   $\Box$ 

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# [Backup] Sample range of o.s. of uniform distribution

$$
f_{X_{(1)},X_{(n)}}(x,y)=\frac{n!}{(n-2)!}[F(y)-F(x)]^{n-2}f(x)f(y)
$$

Let  $X_1, X_2, \cdots, X_N$  be N random variables *i.i.d* with  $X \sim$  Uniform(0, 1)

$$
f_{X_{(1)}, X_{(n)}}(x, y) = n(n-1)[y - x]^{n-2} \text{ where } 0 \le x \le y \le 1
$$

change of variable  $w = y - x$  and  $z = x$ 

$$
f_{X_{(1)},X_{(n)}}(w,z)=n(n-1)w^{n-2}
$$

marginalising,  $y \le 1 \longrightarrow w + z \le 1$ 

$$
f_{X_{(1)},X_{(n)}}(w) = \int_0^{1-w} f_{X_{(1)},X_{(n)}}(w,z)dz
$$
  
=  $n(n-1)w^{n-2}(1-w)$   
= Beta(w; n - 1, 2)

Thus sample range of order statistics of Uniform distribution

$$
f_{1,n}(w) \sim \mathsf{Beta}(w; n-1,2)
$$

**Shifted clustering,**  $r = 0.03$  **Hz,**  $w = 20$  **s,**  $N_{simulations} = 10^7$ 



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**Dynamic Uniform** clustering,  $w = 20$  s,  $\tau_{short} = 10$  ms and  $\tau_{long} = 1$  s



**Dynamic Ramp** clustering,  $w = 20$  s,  $\tau_{short} = 10$  ms and  $\tau_{long} = 1$  s



**Shifted Uniform** clustering,  $w = 20$  s,  $\tau_{short} = 10$  ms and  $\tau_{long} = 1$  s



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<span id="page-48-0"></span>**Shifted Ramp** clustering,  $w = 20$  s,  $\tau_{short} = 10$  ms and  $\tau_{long} = 1$  s

