

Statistical methods for the search for CP violation in $D^0 \rightarrow K_S^0 K^{\mp} \pi^{\pm}$ decays

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Topics:

- CP symmetry and violation
- The $D^0 \rightarrow K_S^0 K^{\mp} \pi^{\pm}$ decay
- The data sample
- The PDF
- The Amplitude Model
- Hypothesis testing (LMP, test statistic, properties)
- Point estimation (MM, estimator, properties)
- Monte Carlo

Disclaimer

This is a simplified version of the actual analysis. The following topics won't be analysed in full details

- RS and WS decay channel difference
- Systematic uncertainties
- Resonances
- Background
- Efficiency
- Parameters cross-talk

CP symmetry

It is defined as the **invariance of fundamental interactions** under the combined symmetry transformations of **charge conjugation (C)** and **parity inversion (P)**

CP violation

Laws of physics are NOT the same if a **particle** is interchanged with the corresponding **antiparticle** while its spatial **coordinates are inverted**

CP symmetry

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CP violation

In short: **particles** and **antiparticles** behave differently

CP symmetry

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CP violation

In short: **particles** and **antiparticles** behave differently

Particle

$$D^0 \rightarrow K_S^0 K^- \pi^+$$

\neq

Antiparticle

$$\bar{D}^0 \rightarrow K_S^0 K^+ \pi^-$$

The $D^0 \rightarrow K_S^0 K^\mp \pi^\pm$ decay channel

Particle

$$D^0 \rightarrow K_S^0 K^- \pi^+$$

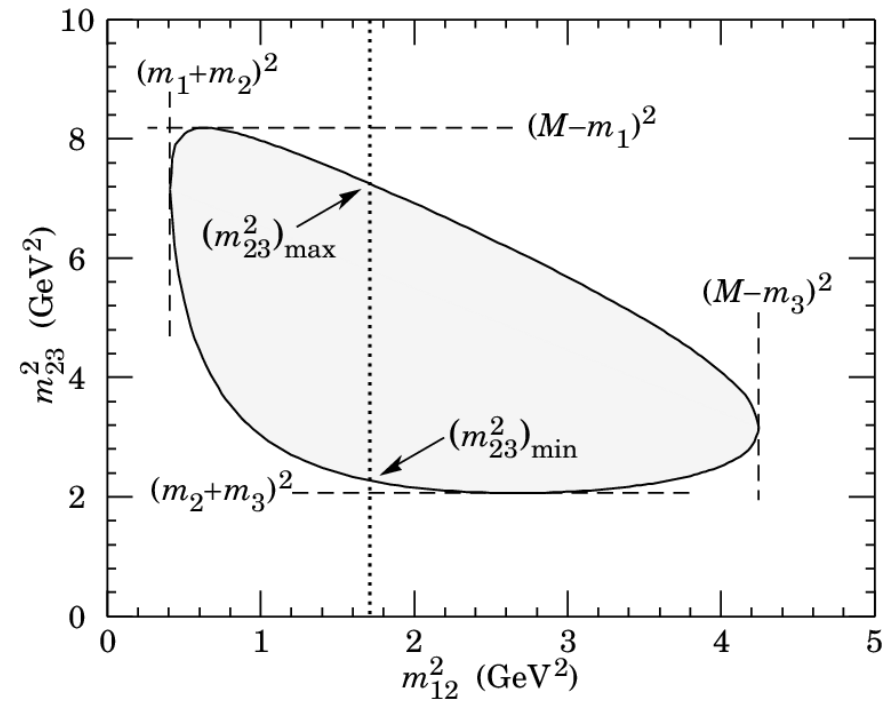
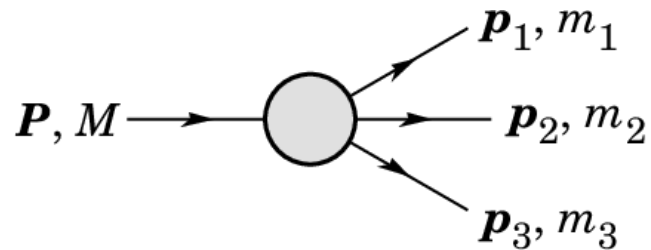
\neq

Antiparticle

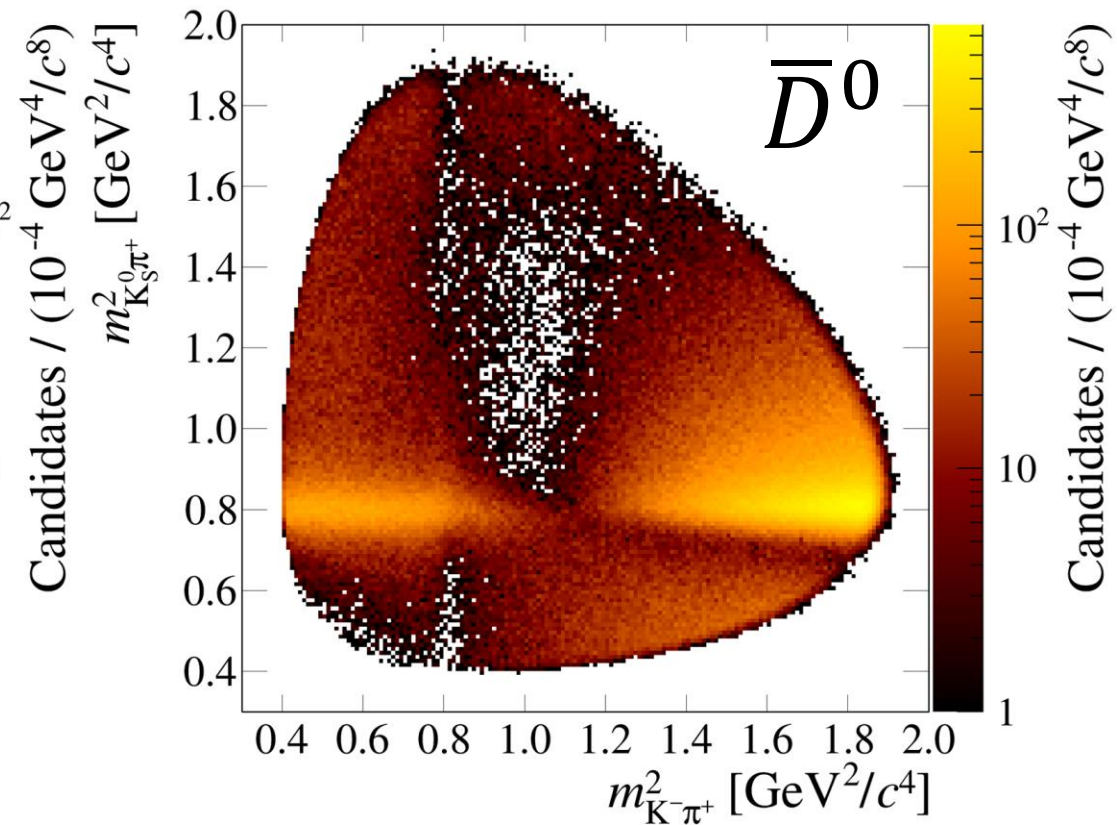
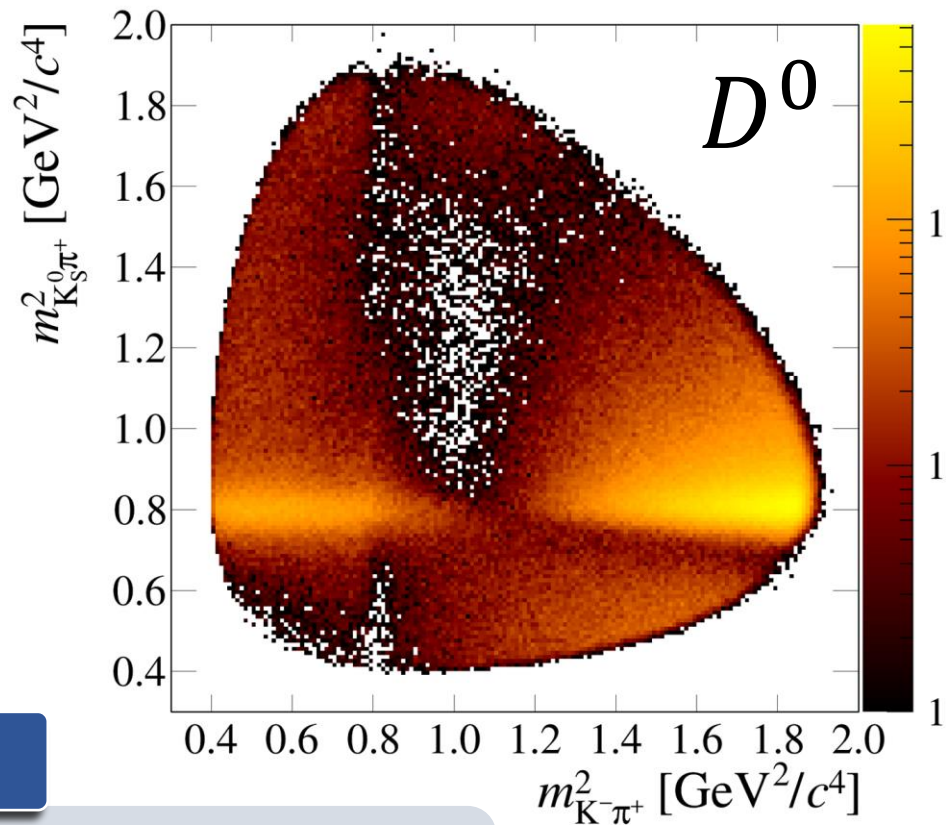
$$\bar{D}^0 \rightarrow K_S^0 K^+ \pi^-$$

Note:

- They are **three-body decays**
- Can be **displayed by a Dalitz Plot**
- May present **resonances**



The data sample

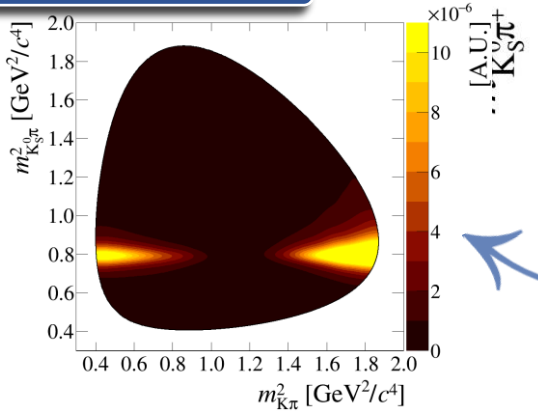


Data sample

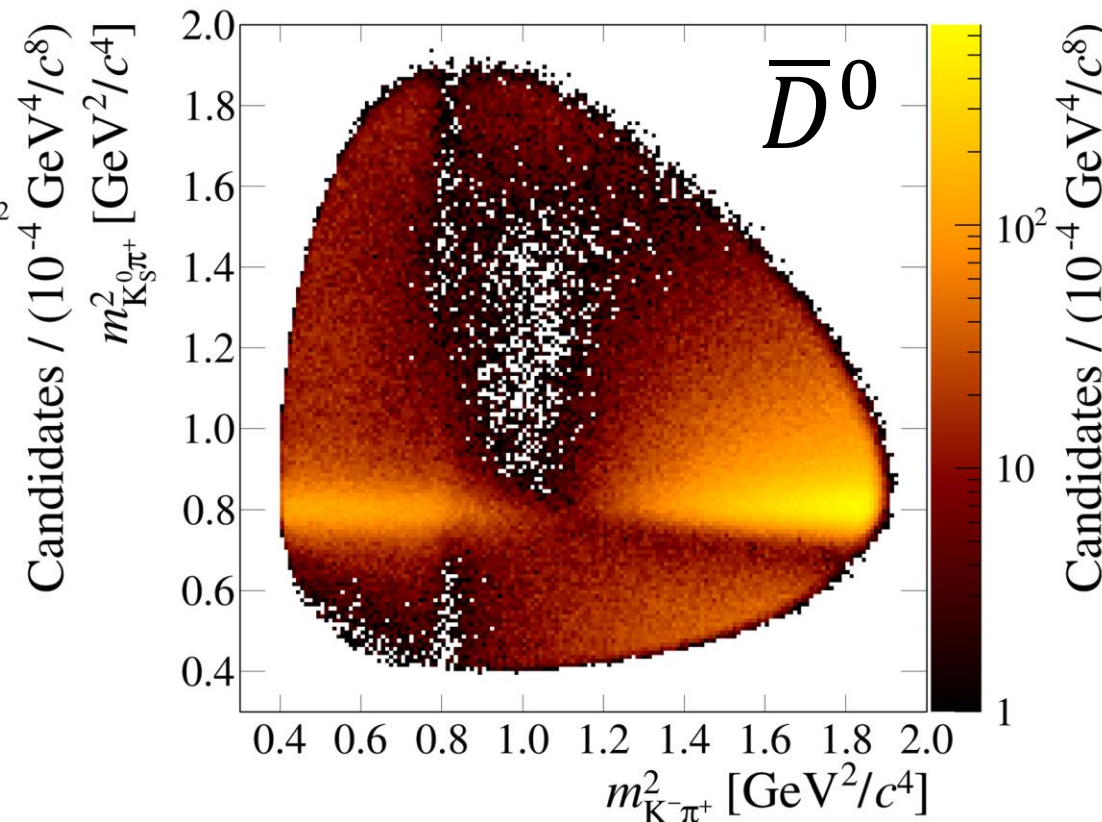
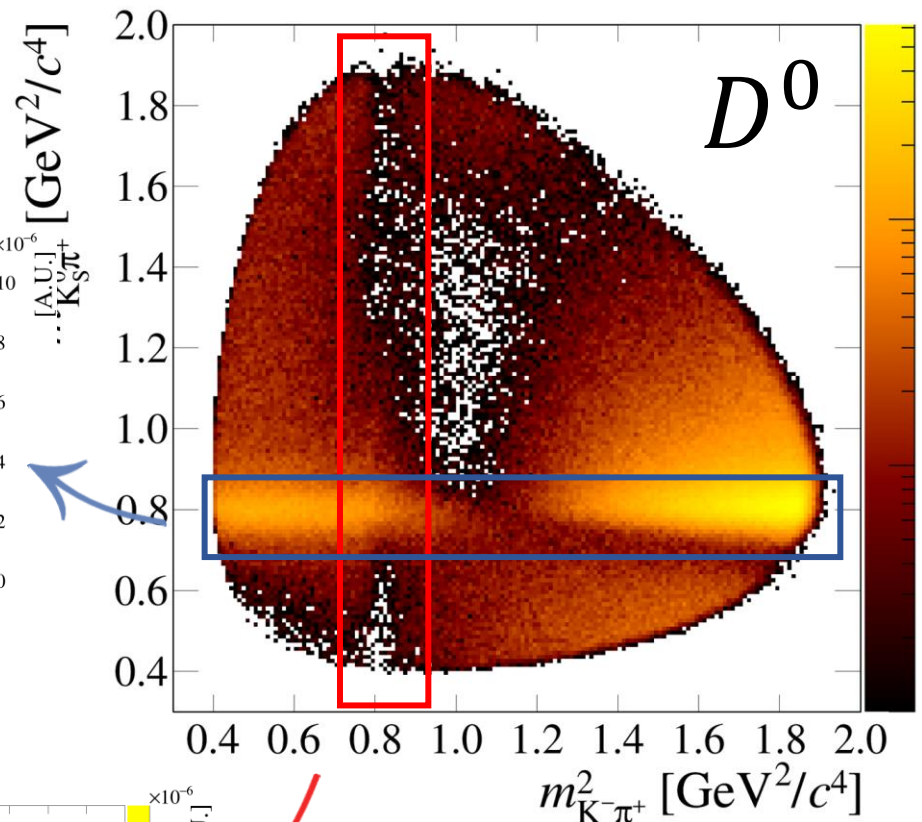
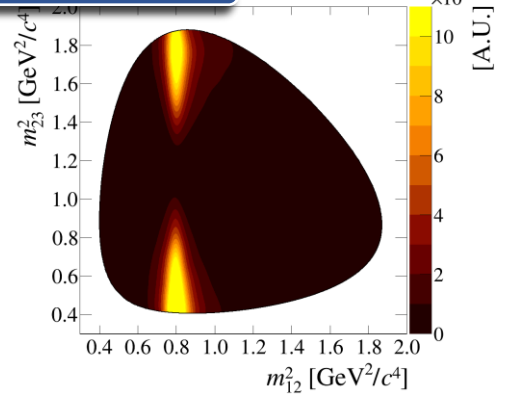
- LHCb Run 2 (2016-2018)
- $N = 900\text{K}$ events
- Background $< 4,5\%$
- **Toy data** (analysis is blinded)
- **Many interfering resonances**

The data sample

$K^*(892)^\pm$

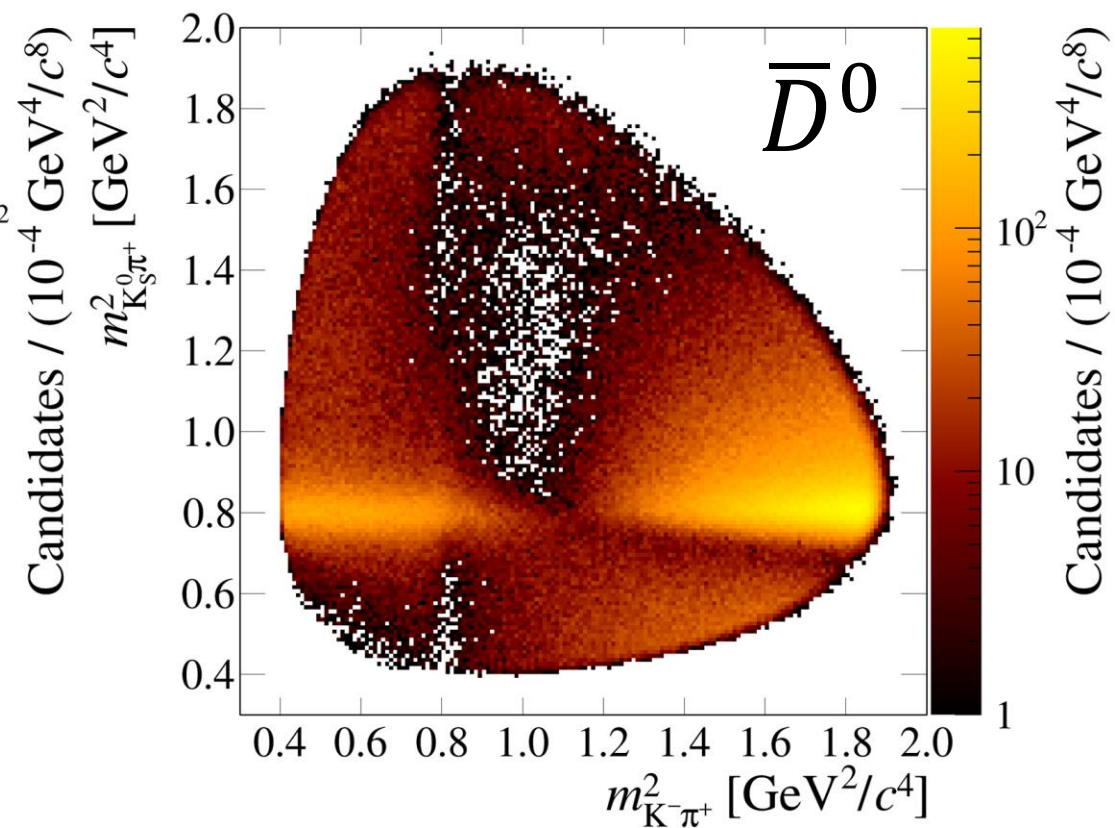
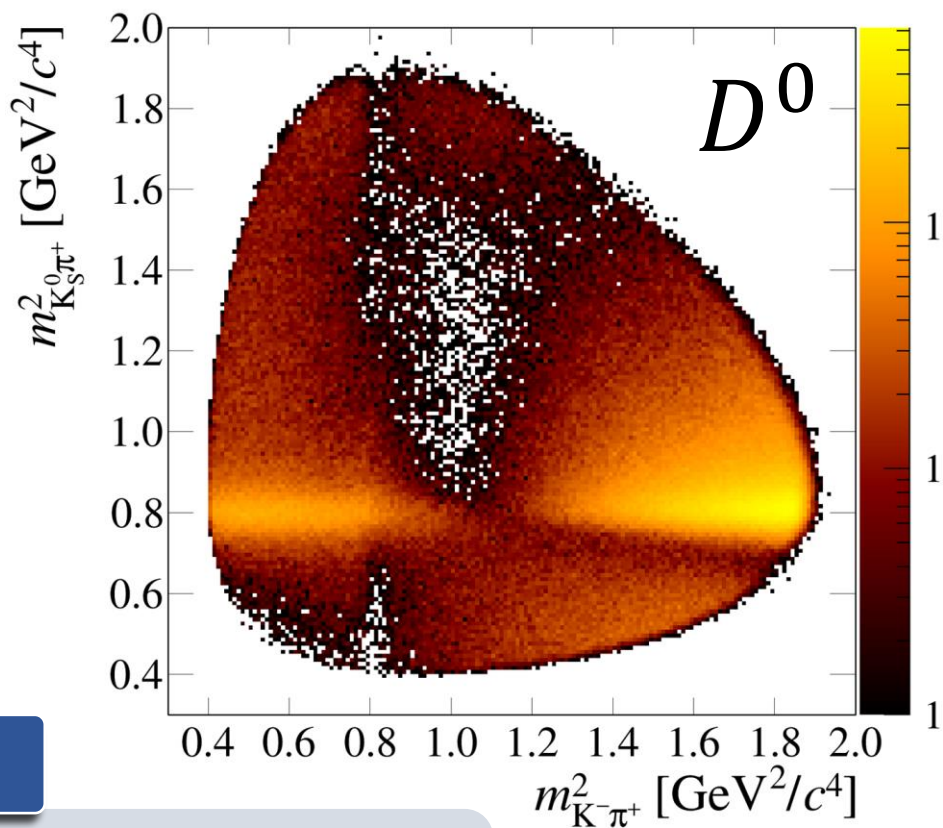


$K^*(892)^0$



$D^0 \rightarrow R(AB) C$		A	B	C
$D^0 \rightarrow K_S^0 (\bar{K}^*)^0$	$(\bar{K}^*)^0 \rightarrow K^\pm \pi^\mp$	π	K	K_S^0
$D^0 \rightarrow K^\mp K^{*\pm}$	$K^{*\pm} \rightarrow K_S^0 \pi^\pm$	K_S^0	π	K
$D^0 \rightarrow \pi^\mp (\rho^\pm, a^\pm)$	$(\rho^\pm, a^\pm) \rightarrow K_S^0 K^\pm$	K	K_S^0	π

The data sample



Data sample

- LHCb Run 2 (2016-2018)
- $N = 900\text{K}$ events
- Background $< 4,5\%$
- **Toy data** (analysis is blinded)
- **Many interfering resonances**

Goal

Search for any **statistically significant differences** between the two data sets that are **related to CP violation**

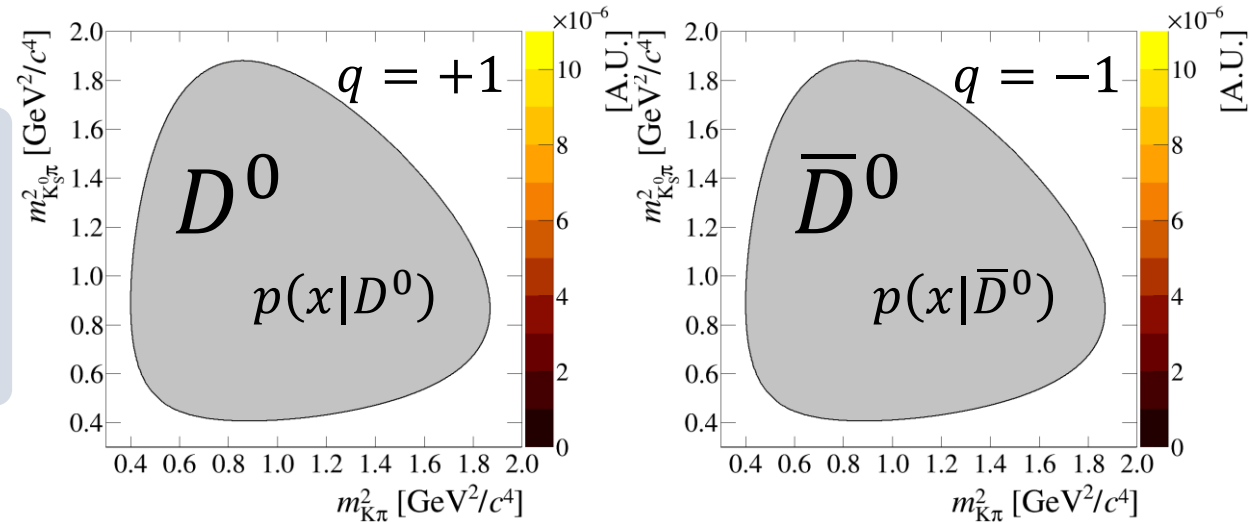
PDF needed

The PDF : derivation

PDF

$$p_{\{\pm 1\}, \Omega}(X) = \dots?$$

$$X = (q, x) \quad \begin{cases} q = \text{flavour } D^0, \bar{D}^0 & q \in \{\pm 1\} \\ x = \text{Dalitz plot position} & x \in \Omega \end{cases}$$



The PDF: derivation

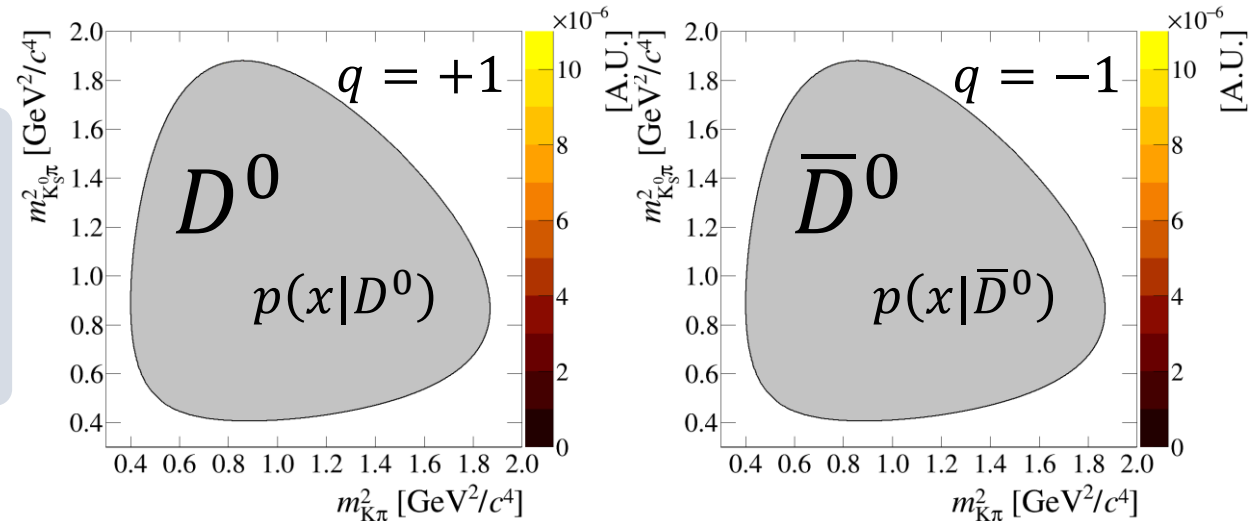
PDF

$$p_{\{\pm 1\},\Omega}(X) = \{\alpha p(x|D^0)\}^{\frac{1+q}{2}} \{(1-\alpha)p(x|\bar{D}^0)\}^{\frac{1-q}{2}}$$

$$X = (q, x) \quad \begin{cases} q = \text{flavour } D^0, \bar{D}^0 & q \in \{\pm 1\} \\ x = \text{Dalitz plot position} & x \in \Omega \end{cases}$$

Proof

$$\begin{aligned} p_{\{\pm 1\},\Omega}(X) &= p_{\{\pm 1\},\Omega}(q, x) = p(q \in \{\pm 1\} \cap x \in \Omega) = \\ &= P_{\{\pm 1\}}(q)p_{\Omega}(x|q) = \\ &= \begin{cases} P(D^0)p(x|D^0) & q = +1 \\ P(\bar{D}^0)p(x|\bar{D}^0) & q = -1 \end{cases} = \\ &= \begin{cases} \alpha p(x|D^0) & q = +1 \\ (1-\alpha)p(x|\bar{D}^0) & q = -1 \end{cases} = \\ &= \{\alpha p(x|D^0)\}^{\frac{1+q}{2}} \{(1-\alpha)p(x|\bar{D}^0)\}^{\frac{1-q}{2}} \end{aligned}$$

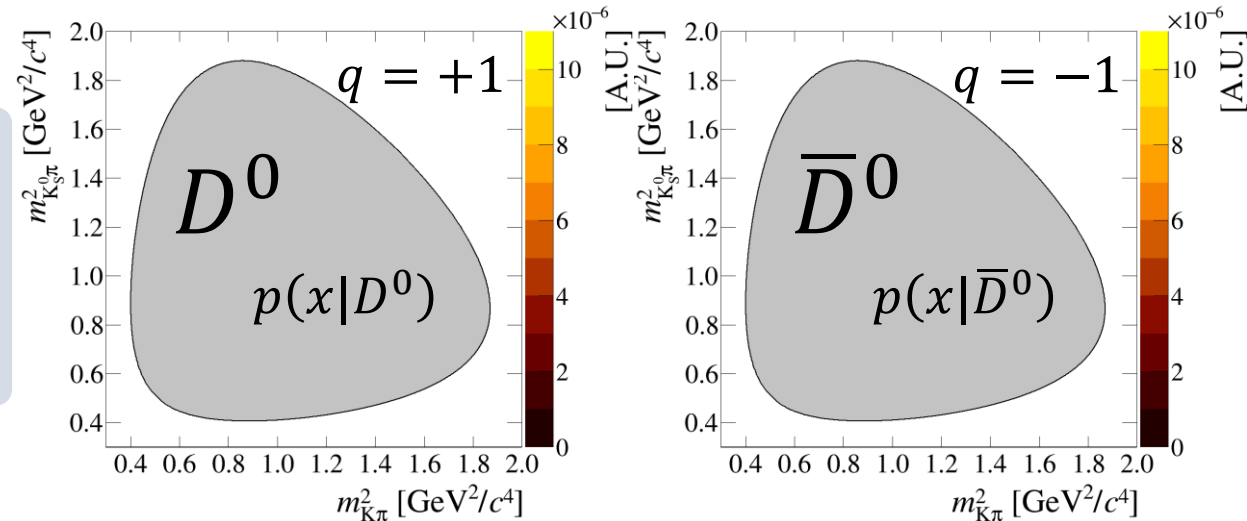


- $P(A, B) = P(A \cap B)$ joint probability
- $P(B|A) = \frac{P(A, B)}{P(A)}$ conditional probability
- $q = +1 \rightarrow D^0$, $q = -1 \rightarrow \bar{D}^0$
- $P(D^0) = \alpha$, $P(\bar{D}^0) = 1 - \alpha$
 $\alpha = \text{global asymmetry}$ $\alpha \in [0, 1]$
- Rearrange

PDF

$$p_{\{\pm 1\}, \Omega}(X) = \{\alpha p(x|D^0)\}^{\frac{1+q}{2}} \{(1-\alpha)p(x|\bar{D}^0)\}^{\frac{1-q}{2}}$$

$$X = (q, x) \quad \begin{cases} q = \text{flavour } D^0, \bar{D}^0 & q \in \{\pm 1\} \\ x = \text{Dalitz plot position} & x \in \Omega \end{cases}$$



α

The **red terms** can be interpreted as a **Bernoulli distribution** where α is the success probability

$$f(q; \alpha) = \alpha^{\frac{1+q}{2}} (1-\alpha)^{\frac{1-q}{2}}$$

$$P(D^0) = \alpha \quad P(\bar{D}^0) = 1 - \alpha$$

$p(x|q)$

The **green terms** are the **flavour specific PDF** of the data Dalitz plot distribution

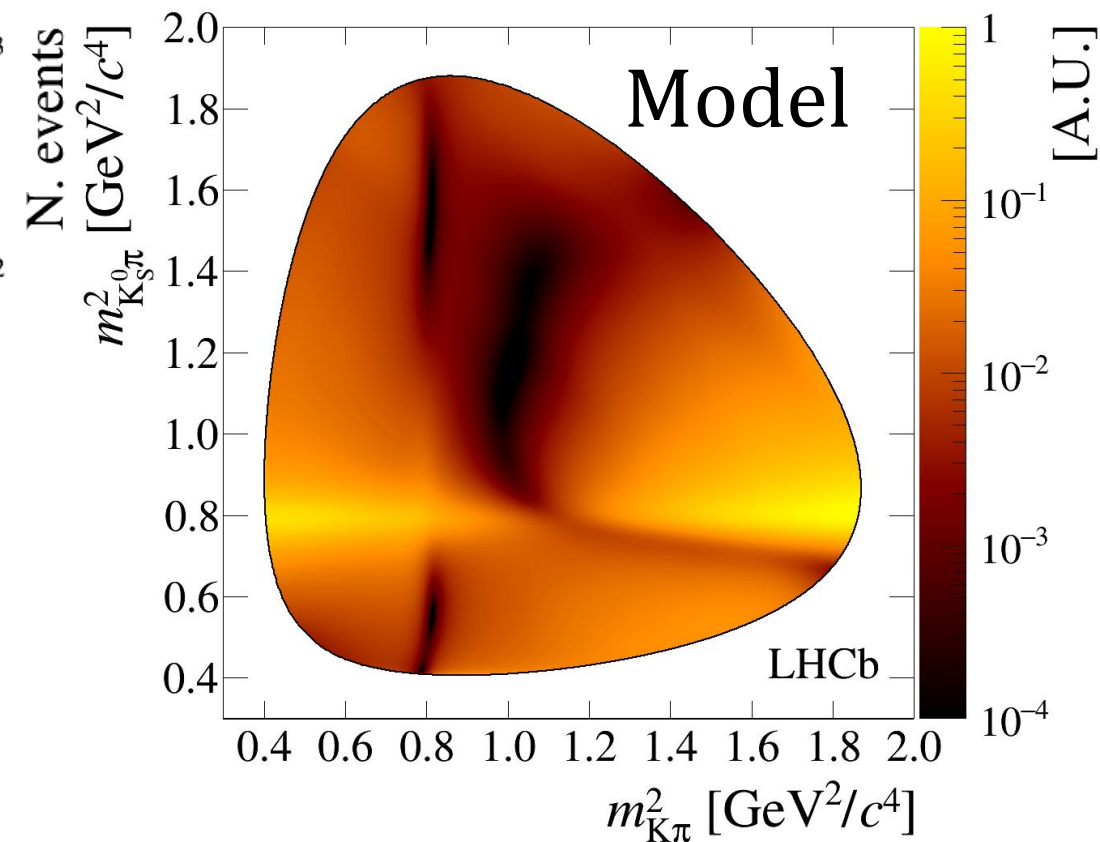
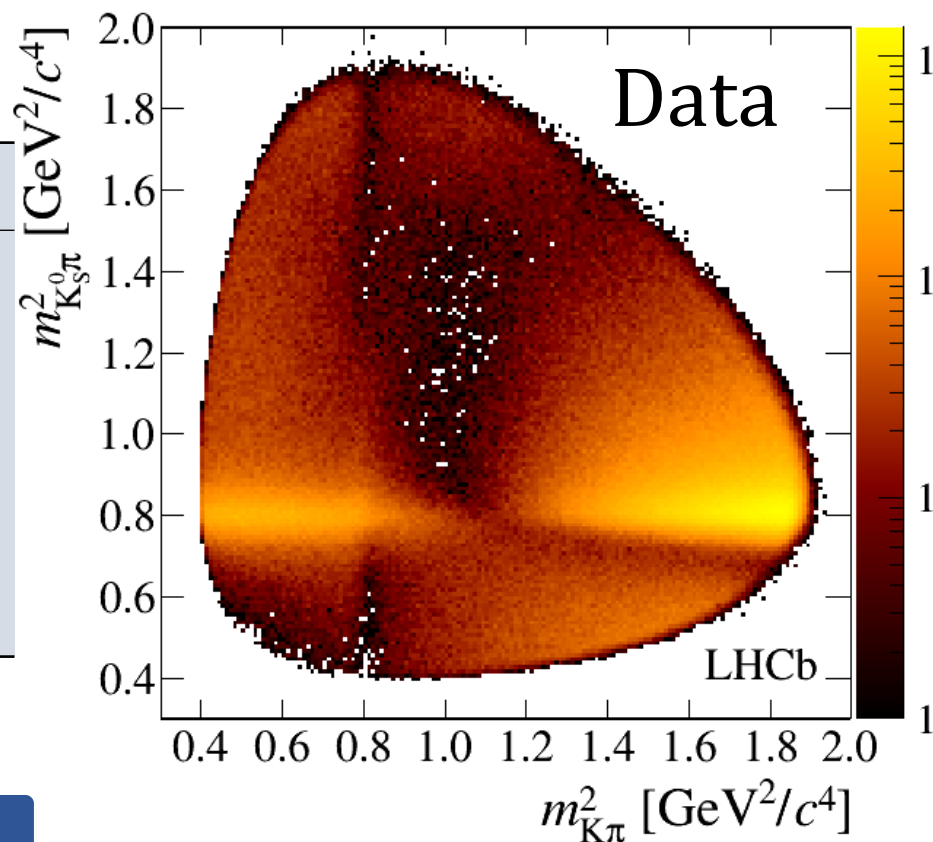
$$p_{\Omega}(x|q = +1) = p(x|D^0)$$

$$p_{\Omega}(x|q = -1) = p(x|\bar{D}^0)$$

To be defined 

The PDF: amplitude model from LHCb Run 1

$R \rightarrow K\pi$	$R \rightarrow K_S^0 K^\pm$
$K^*(892)^0$	$a_0(892)^-$
$K^*(892)^\pm$	$a_2(1320)^+$
$K_1^*(1410)^{0,\pm}$	$a_0(1450)^\pm$
$K_0^*(1430)^{0,\pm}$	$\rho(1450)^\pm$
$K_2^*(1430)^0$	$\rho(1700)^\pm$



Isobar model

$$A_{K_S^0 K^\mp \pi^\pm}(x) = \varepsilon(x) \left| \sum_R a_R e^{i\phi_R} \mathcal{M}_R(x) \right|^2$$

- $x = (m_{K\pi}^2, m_{K_S^0\pi}^2)$ position in the Dalitz plot
- The sum runs over 2-body intermediate resonances R
- $\mathcal{M}_R(x)$ is the matrix element, $\varepsilon(x)$ is the efficiency model

LHCb collaboration, R. Aaij et al., "Studies of the resonance structure in $D^0 \rightarrow K_S^0 K^\mp \pi^\pm$ decays", [[PRD.93.052018](#)]

CPV parameters

The amplitude model can be re-parametrized using

$$p_{\Omega}(x|q) \equiv p_{\pm}(x; \vec{\theta}) = \varepsilon(x) \left| \sum_R a_R (1 \pm \Delta \mathbf{a}_R) e^{i(\phi_R \pm \Delta \phi_R)} \mathcal{M}_R(x) \right|^2 \quad \begin{cases} D^0 & \rightarrow + \\ \bar{D}^0 & \rightarrow - \end{cases}$$

So, $\vec{\theta} = (\Delta \mathbf{a}, \Delta \phi)_R$ are the set of **CP -violating parameters**.

- A search for CPV must deal with these $2N_R$ parameters
- A_{CP} is a function of $\vec{\theta} = (\Delta \mathbf{a}, \Delta \phi)_R$

Linear approximation

Since **CPV is expected to be small**, the $p_{\pm}(x; \vec{\theta})$ can be expanded in $\vec{\theta} = \vec{0}$ as

$$p_{\pm}(x; \vec{\theta}) = f(x) \pm \sum_{\chi \in \vec{\theta}} \chi g_{\chi}(x) + O(\vec{\theta}^2)$$

where

$$f(x) \equiv p_{\pm}(x; \vec{\theta}) \Big|_{\vec{\theta}=\vec{0}} = p(x; \vec{0}) \quad g_{\chi}(x) \equiv \frac{\partial}{\partial \chi} p_{\pm}(x, \vec{\theta}) \Big|_{\vec{\theta}=\vec{0}}$$

PDF

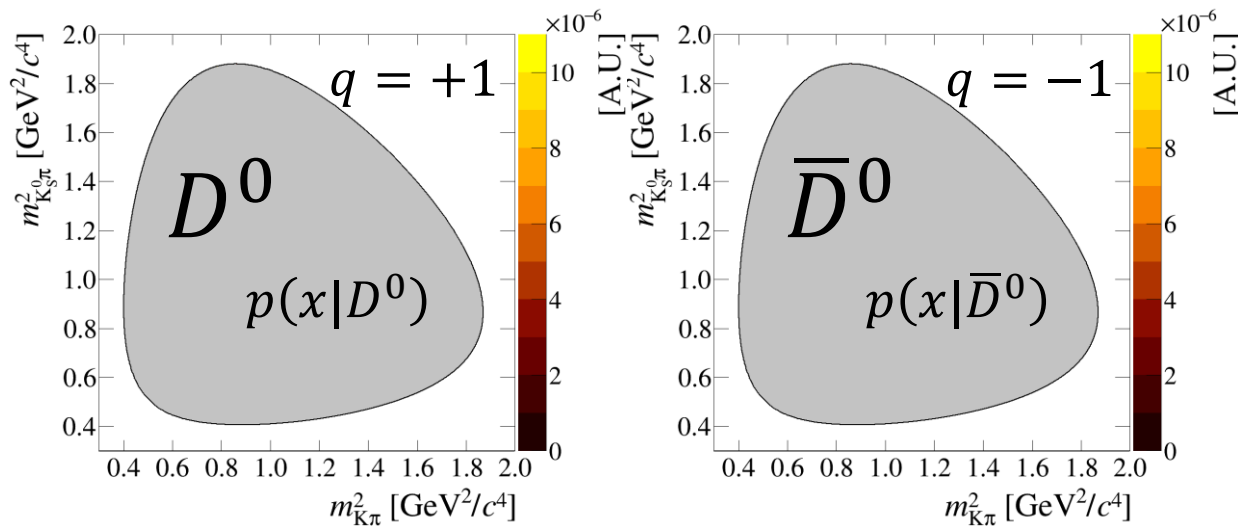
$$p(q, x; \alpha, \vec{\theta}) = \{\alpha p_+(x; \vec{\theta})\}^{\frac{1+q}{2}} \{(1 - \alpha)p_-(x; \vec{\theta})\}^{\frac{1-q}{2}}$$

- $q = \text{flavour } D^0, \bar{D}^0$ and $q \in \{\pm 1\}$
- $x = \text{Dalitz plot position}$ and $x \in \Omega$
- $\alpha = \text{global asymmetry}$ and $\alpha \in [0, 1]$
- $p_{\pm}(x; \vec{\theta}) = f(x) \pm \sum_{\chi \in \vec{\theta}} \chi g_{\chi}(x)$ linear approximation

Goal

The search for CP violation may be carried out through:

- Hypothesis testing
- Point estimation of $\vec{\theta} = (\Delta a, \Delta \phi)_R$



Hypothesis test

Hp. test

The search for CPV can be interpreted as a **hypothesis test**

- $H_0 = CP$ conservation $\vec{\theta} = \vec{0}$
- $H_1 = CP$ violation $\vec{\theta} \neq \vec{0}$

Considering that:

- H_1 **depends on $\vec{\theta}$** and is not simple
- Predictions say $CPV \sim 0$

the most convenient test is the **Locally Most Powerful (LMP)** test

LMP test

- It can be used when $H_0 \sim H_1$
- It is the **most powerful test** $\forall \vec{\theta} \in I(\vec{\theta}_0)$
- For a data set \vec{X} of N observations, **the test statistic** is

$$t_{\chi}(\vec{X}) = s(\chi) \Big|_{\vec{\theta}=\vec{0}} = \frac{\partial}{\partial \chi} \log \mathcal{L}_{\vec{X}}(\vec{\theta}) \Big|_{\vec{\theta}=\vec{0}} \quad \forall \chi \in \vec{\theta}$$

Likelihood

$$\mathcal{L}_X(\vec{\theta}, \alpha) = \{\alpha \mathcal{L}_+(\vec{\theta}|x)\}^{\frac{1+q}{2}} \{(1-\alpha) \mathcal{L}_-(\vec{\theta}|x)\}^{\frac{1-q}{2}}$$

- $\mathcal{L}_\pm(\vec{\theta}|x) = f(x) \pm \sum_{\chi \in \vec{\theta}} \chi g_\chi(x)$ linear approximation
- $\vec{\theta} \in I(\vec{\theta}_0 = \vec{0})$ and $\alpha \in [0,1]$

$t_\chi(\vec{X})$

$$\begin{aligned} t_\chi(\vec{X}) &= s(\chi) \Big|_{\vec{\theta}=\vec{0}} = \frac{\partial}{\partial \chi} \log \mathcal{L}_{\vec{X}}(\vec{\theta}) \Big|_{\vec{\theta}=\vec{0}} = \\ &= \frac{\partial}{\partial \chi} \ln \prod_{i=1}^N [\alpha \mathcal{L}_+(\vec{\theta}|x_i)]^{\frac{1+q_i}{2}} [(1-\alpha) \mathcal{L}_-(\vec{\theta}|x_i)]^{\frac{1-q_i}{2}} \Big|_{\vec{\theta}=\vec{0}} = \\ &= \sum_{i=1}^N \left(\frac{1+q_i}{2} \right) \frac{g_\chi(x_i)}{f(x_i) + \sum_{\psi \in \vec{\theta}} \psi g_\psi(x_i)} - \left(\frac{1-q_i}{2} \right) \frac{g_\chi(x_i)}{f(x_i) - \sum_{\psi \in \vec{\theta}} \psi g_\psi(x_i)} \Big|_{\vec{\theta}=\vec{0}} = \\ &= \sum_{i=1}^N q_i \frac{g_\chi(x_i)}{f(x_i)} \quad \forall \chi \in \vec{\theta} \end{aligned}$$

Hypothesis testing: properties

$$t_\chi(\vec{X})$$

The test statistic $t_\chi(\vec{X})$ is

$$t_\chi(\vec{X}) = \sum_{i=1}^N q_i \frac{g_\chi(x_i)}{f(x_i)}$$

$$t_{thr}(\alpha)$$

At fixed α , the threshold $t_{thr}(\alpha)$ is given by

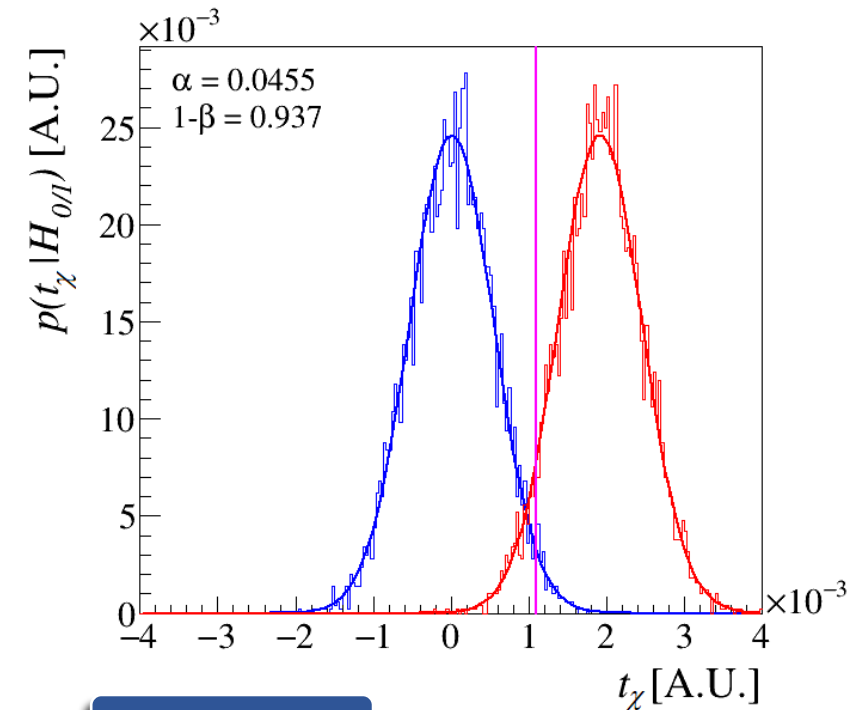
$$\int_{t_{thr}(\alpha)}^{+\infty} p(t_\chi|H_0) dt_\chi = \alpha$$

α is the
significance
level

$$\text{pow}(\chi)$$

At given χ , the power $\text{pow}(\chi)$ is given by

$$\text{pow}(\chi) = \int_{t_{thr}(\alpha)}^{+\infty} p(t_\chi|H_1) dt_\chi$$



Example

- $\chi = \Delta\phi_{K^*(892)^0}$
- $p(t_\chi|H_0) \rightarrow \chi_0 = 0^\circ$
- $p(t_\chi|H_1) \rightarrow \chi_1 = 1.5^\circ$

Point estimation

Method of Moments

MLE vs MM

Given the complexity of the $p(q, x; \alpha, \vec{\theta})$

- the Maximum Likelihood Estimator (**MLE**) is **unfeasible**
- the Method of Moments (**MM**) could be **useful**

In its most general form, **the MM states that**

if $\forall \chi \in \vec{\theta}$ $\left\{ \begin{array}{l} \exists \text{Var}[t_\chi(\mathbf{x})] \text{ and it is finite} \\ h_\chi(\vec{\theta}) \equiv \mathbf{E}[t_\chi(\mathbf{x})] \\ \exists h_\chi^{-1} \text{ inverse of } h_\chi(\vec{\theta}) \text{ at least } \forall \vec{\theta} \in I(\vec{\theta}_0) \end{array} \right.$

then any $\chi \in \vec{\theta}$ can be estimated using the **estimator** defined as

$$\hat{t}_\chi(\vec{X}) = h^{-1} \left(\frac{1}{N} \sum_{i=1}^N t_\chi(X_i) \right)$$

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LMP statistic

Let's look at the **properties** of the **LMP statistic**

- $t_\chi(X) = q \frac{g_\chi(x)}{f(x)}$
- $\mathbf{E}[t_\chi] = A_{\chi\psi} \theta_\psi$
- $\text{Cov}[t_\chi, t_\psi] = A_{\chi\psi} + (2\alpha - 1) B_{\chi\psi\xi} - \mathbf{E}[t_\chi] \mathbf{E}[t_\psi]$

MM is satisfied

The coefficients are:

$$A_{\chi\psi} \equiv \int_{\Omega} \frac{g_\chi(x) g_\psi(x)}{f(x)} dx$$

$$B_{\chi\psi\xi} \equiv \int_{\Omega} \frac{g_\chi(x) g_\psi(x)}{f(x)} dx$$

Estimator

Defined $\mathbb{A} \equiv (A_{\chi\psi})_{\chi,\psi \in \vec{\theta}}$ and $\mathbb{B} \equiv (B_{\chi\psi\xi})_{\chi,\psi,\xi \in \vec{\theta}}$, the **MM estimator** is given by

$$\hat{\mathbf{t}}_{\chi}(\vec{\mathbf{X}}) = \frac{1}{N} \mathbb{A}_{\chi\psi}^{-1} \sum_{i=1}^N q_i \frac{\mathbf{g}_{\psi}(\mathbf{x}_i)}{f(\mathbf{x}_i)}$$

- $E[\hat{t}_{\chi}] = \theta_{\chi}$
- $\text{Cov}[\hat{t}_{\chi}, \hat{t}_{\psi}] = \frac{1}{N} \mathbb{A}_{\chi\psi}^{-1} [\mathbb{A}_{\chi\psi} + (2\alpha - 1) \mathbb{B}_{\chi\psi\xi} \theta_{\xi}] (\mathbb{A}_{\chi\psi}^{-1})^T - \frac{1}{N} \theta_{\chi} \theta_{\psi}$

Properties

The estimator is

- Asymptotically normally distributed
- Consistent
- Unbiased
- Highly efficient ($\varepsilon > 91\%$)
- Properties proved analytically

Good Estimator!

Point estimation: Monte Carlo simulation

Method

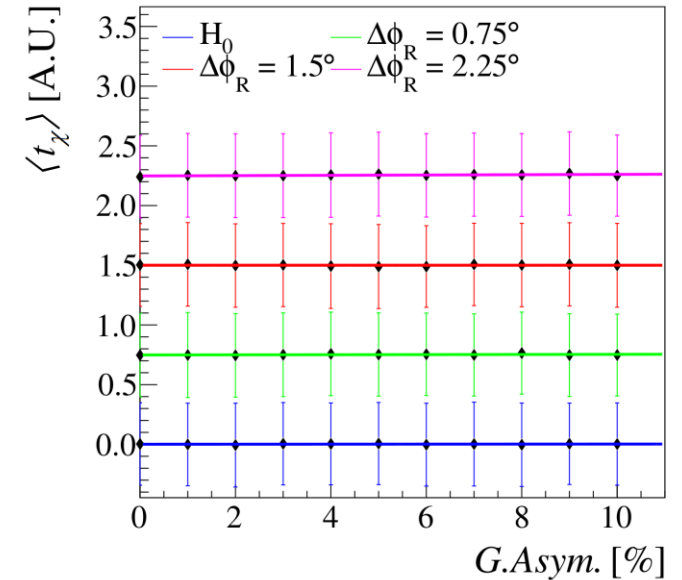
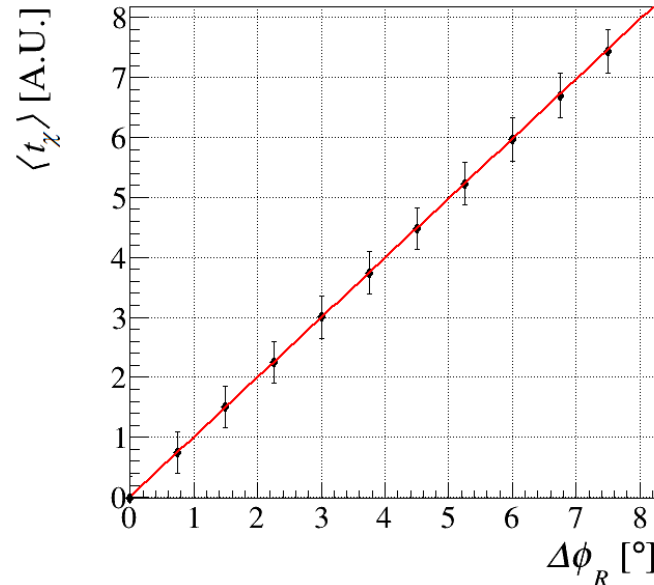
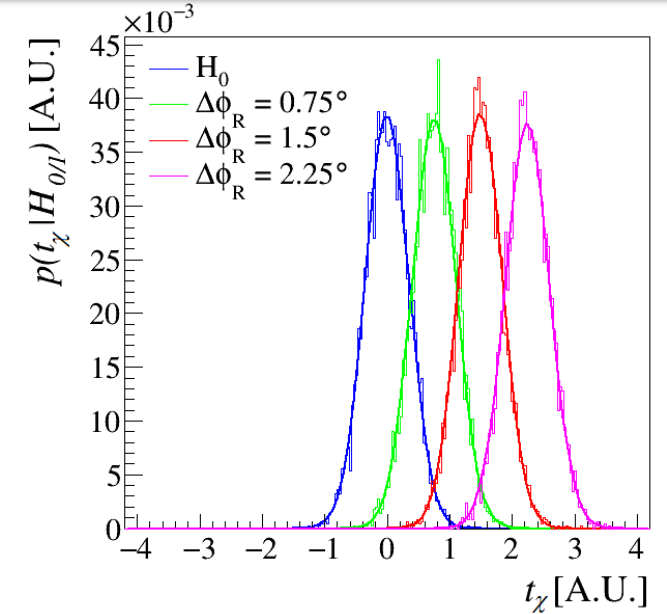
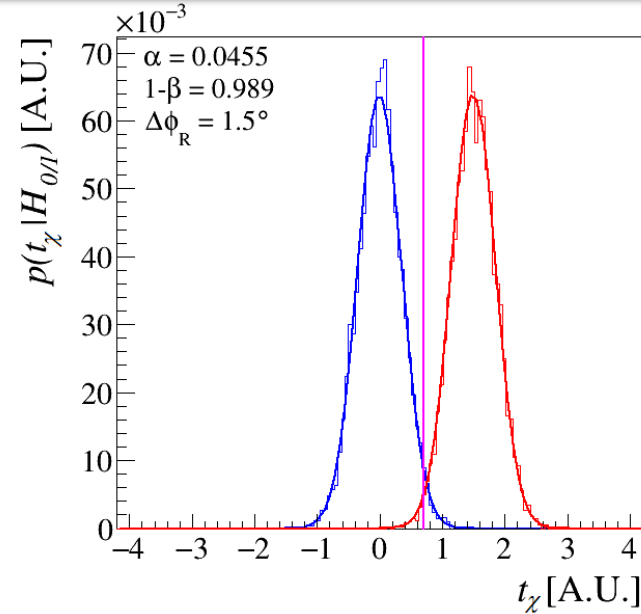
1. Inject CPV in the model through a selected $\chi \in \vec{\theta}$
2. Generate two random samples $S_+(D^0)$ and $S_-(\bar{D}^0)$
3. Evaluate \hat{t}_θ using its definition

$$\hat{t}_\chi(\vec{X}) = \frac{1}{N} A_{\chi\psi}^{-1} \sum_{i=1}^N q_i \frac{g_\psi(x_i)}{f(x_i)}$$

4. Repeat 5000 times to obtain $p(\hat{t}_\chi; \alpha, \vec{\theta})$
5. Vary χ and α over the expected theoretical and experimental ranges

Example

- $\chi = \Delta\phi_{K^*(892)^0}$
- Same results for other parameters



Point estimation: Monte Carlo Simulation

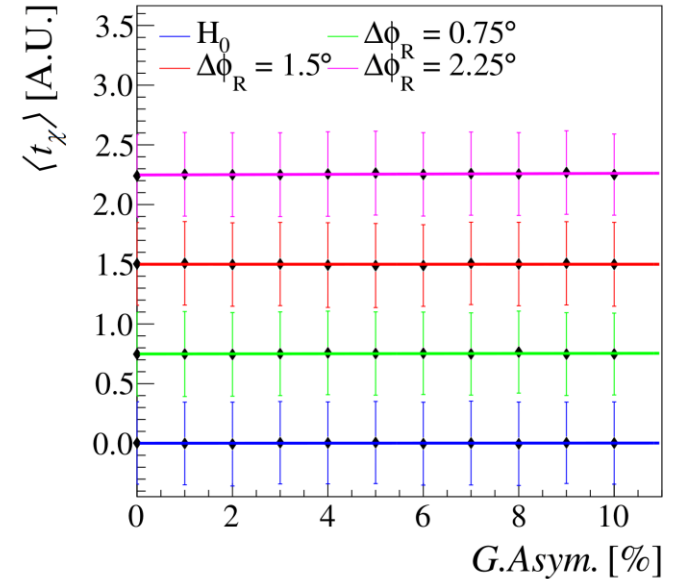
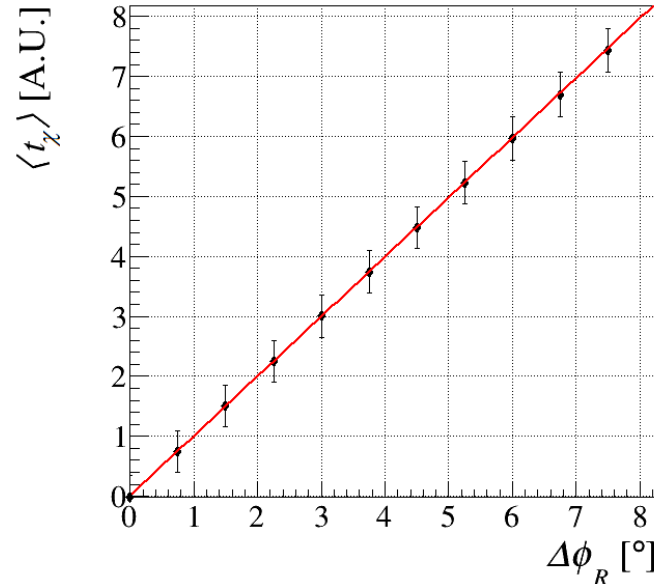
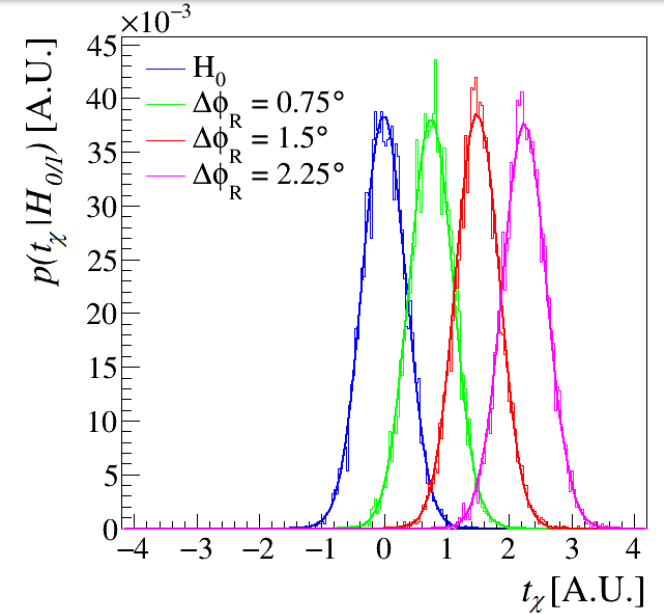
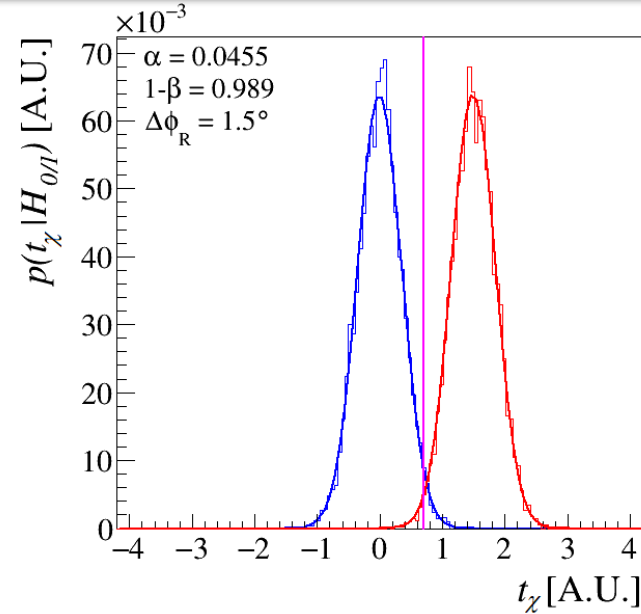
Results

The numerical analysis validates the estimator's properties:

- The **asymptotic normality** of $p(\hat{t}_\chi; \alpha, \vec{\theta})$ implies \hat{t}_χ **consistency**
- The **linear** fit coincides with the bisector. This implies \hat{t}_χ **unbiasedness**
- The variance obtained from the fit is equal to the analytical prediction. The \hat{t}_χ **high efficiency** is confirmed
- The $E[\hat{t}_\chi]$ shows **no dependence** from the global asymmetry α

Example

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- Same results for other parameters



**Thank you for your
attention**